

QUEEN MARY UNIVERSITY OF LONDON

MTH5120

Statistical Modelling I

Solution to Exercise Sheet 8

1. Based on the Boston dataset available on the library MASS, relative to Housing Values in Suburbs of Boston. The variables of interest are:

- Y equal to *medv* is median value of owner-occupied homes in \$1000.
- X_1 equal to *lstat* is the lower status of the population (percent)
- X_2 equal to *rm* is the average number of rooms per dwelling
- X_3 equal to *age* is the proportion of owner-occupied units built prior to 1940

- (a) We fit the Model 1: $Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 x_{3i} + \varepsilon_i$, where $\varepsilon_i \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$.

We use the following R commands for loading the data:

```
> library(MASS)
> data("Boston")
> attach(Boston)
```

The following objects are masked from Boston (pos = 13):

```
age, black, chas, crim, dis, indus, lstat, medv, nox,
ptratio, rad, rm, tax, zn
```

Then we fit the Model 1 to the data:

```
> fitlm1 <- lm(medv ~ lstat + rm + age)
> summary(fitlm1)
```

Call:

```
lm(formula = medv ~ lstat + rm + age)
```

Residuals:

```
      Min       1Q   Median       3Q      Max
-18.210  -3.467  -1.053   1.957  27.500
```

Coefficients:

```
              Estimate Std. Error t value Pr(>|t|)
(Intercept) -1.175311    3.181924  -0.369    0.712
lstat        -0.668513    0.054357 -12.298 <2e-16 ***
rm           5.019133    0.454306  11.048 <2e-16 ***
age          0.009091    0.011215   0.811    0.418
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Residual standard error: 5.542 on 502 degrees of freedom

Multiple R-squared: 0.639, Adjusted R-squared: 0.6369

F-statistic: 296.2 on 3 and 502 DF, p-value: < 2.2e-16

- (b) Looking at the last line of the command summary, we find that the F-Test is equal to 296.2 and there is strong evidence against the null hypothesis $H_0 : \beta_1 = \beta_2 = \beta_3 = 0$ and the R^2 is equal to 63% similar to the adjusted R^2 .
- (c) Moving to the parameters of interest, we look at the summary described above. In this case, we have that there is evidence to reject the null hypothesis $H_0 : \beta_j = 0$ against the alternative $H_1 : \beta_j \neq 0$ for β_1 and β_2 , thus the coefficients for *lstat* and *rm* are statistically significant. On the other hand, the intercept and the parameter related to *age* could not reject the null hypothesis, thus the two coefficients are not statistically significant.
- (d) Moving to the second model, we fit the following Model 2: $Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \varepsilon_i$, where $\varepsilon_i \underset{iid}{\sim} \mathcal{N}(0, \sigma^2)$. By using the R command, we have:

```
> fitlm2 <- lm(medv ~ lstat + rm)
> summary(fitlm2)
```

Call:

```
lm(formula = medv ~ lstat + rm)
```

Residuals:

Min	1Q	Median	3Q	Max
-18.076	-3.516	-1.010	1.909	28.131

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.35827	3.17283	-0.428	0.669
lstat	-0.64236	0.04373	-14.689	<2e-16 ***
rm	5.09479	0.44447	11.463	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 5.54 on 503 degrees of freedom

Multiple R-squared: 0.6386, Adjusted R-squared: 0.6371

F-statistic: 444.3 on 2 and 503 DF, p-value: < 2.2e-16

As previously described, we have that the F-statistic is 444.3, thus the overall regression is statistically significant and there is strong evidence against the null hypothesis. Moving to the parameters, in this scenario the parameter of *lstat* and *rm* are statistically significant, while the intercept continuously remains not statistically significant.

- (e) Regarding the best model, we compare the adjusted R^2 for both the models. For Model 1, $adj(R^2) = 0.6369$, while for Model 2, $adj(R^2) = 0.6371$, thus the Model 2 is the best model and in this case also all the parameters except the intercept are statistically significant.

2. Coursework component

When fitting the model

$$E[Y_i] = \beta_0 + \beta_1 x_{1,i} + \beta_2 x_{2,i}$$

to a set of $n = 25$ observations, the following results were obtained using the general linear model notation:

$$\mathbf{X}^t \mathbf{X} = \begin{pmatrix} 25 & 219 & 10232 \\ 219 & 3055 & 133899 \\ 10232 & 133899 & 6725688 \end{pmatrix}, \quad \mathbf{X}^t \mathbf{Y} = \begin{pmatrix} 559.60 \\ 7375.44 \\ 337071.69 \end{pmatrix}$$
$$(\mathbf{X}^t \mathbf{X})^{-1} = \begin{pmatrix} 0.11321519 & -0.00444859 & -0.000083673 \\ -0.00444859 & 0.00274378 & -0.000047857 \\ -0.00008367 & -0.00004786 & 0.000001229 \end{pmatrix}$$

Also $\mathbf{Y}^t \mathbf{Y} = 18310.63$ and $\bar{Y} = 22.384$.

(a) We find the least square estimator by using

$$\begin{aligned} \hat{\boldsymbol{\beta}} &= (\mathbf{X}^t \mathbf{X})^{-1} \mathbf{X}^t \mathbf{Y} \\ &= \begin{pmatrix} 25 & 219 & 10232 \\ 219 & 3055 & 133899 \\ 10232 & 133899 & 6725688 \end{pmatrix}^{-1} \begin{pmatrix} 559.60 \\ 7375.44 \\ 337071.69 \end{pmatrix} \\ &= \begin{pmatrix} 2.34123 \\ 1.61591 \\ 0.01438 \end{pmatrix} \end{aligned}$$

Thus the related fitted model can be written as

$$y = 2.34123 + 1.61591x_1 + 0.01438x_2$$

(b) Based on the previous results, we can construct the ANOVA table. First of all, we need to define

$$\begin{aligned} SS_R &= \hat{\boldsymbol{\beta}}^t \mathbf{X}^t \mathbf{Y} - n\bar{y}^2 = (2.34123 \quad 1.61591 \quad 0.01438) \cdot \begin{pmatrix} 559.60 \\ 7375.44 \\ 337071.69 \end{pmatrix} - 25 \cdot 22.384^2 \\ &= 18076.9 - 12526.09 = 5550.81 \end{aligned}$$

Moving to the SS_T , we have that

$$SS_T = \mathbf{Y}^t \mathbf{Y} - n\bar{y}^2 = 18310.63 - 12526.09 = 5784.54$$

Thus, we have that $SS_E = SS_T - SS_R = 5784.54 - 5550.81 = 233.73$. Moving to S^2 or the so called MS_E , we have

$$S^2 = \frac{SS_E}{(25 - 3)} = \frac{233.73}{22} = 10.62409$$

Analogously, we have the MS_R , which is

$$MS_R = \frac{SS_R}{25 - 23} = \frac{5550.81}{2} = 2775.405$$

Finally the F statistic, which is

$$F = \frac{MS_R}{MS_E} = \frac{2775.405}{10.62409} = 261.237$$

In conclusion, we can build up the ANOVA table as

Source	df	SS	MS	F	p-value
Regression	2	5550.81	2775.405	261.237	0.0000
Residual	22	233.73	10.62409		
Total	24	5784.54			

3. Based on the previous results:

- We see from the ANOVA table that the value of F is 261.237. The critical value at the 5% significance level for an F_{22}^2 distribution is 3.44 so we can reject the null hypothesis at the 5% significance level.
- Moving to the 95% confidence interval for the three parameters of the model. We start with β_0 and we remind that the variance of β_j is

$$S^2 \cdot c_{jj}$$

dove c_{jj} is the j th diagonal element of $(\mathbf{X}^t \mathbf{X})^{-1}$. Remind in this case that $t_{22}(0.025) = 2.074$.

Thus, $\hat{\beta}_0 = 2.34123$ and then its variance is $S^2 \cdot 0.11321519 = 10.62409 \cdot 0.11321519 = 1.20280$ and then its standard error is the square root of the variance, $\sqrt{1.20280} = 1.096722$.

Thus for β_0 the 95% confidence interval is

$$2.34123 \pm 2.074 \cdot 1.096722 = 2.34123 \pm 2.274601 = (0.066629, 4.615831)$$

Analogously, we can run the confidence interval for the other two coefficients β_1 and β_2 . For β_1 , we have that the standard error is equal to $\sqrt{S^2 \cdot 0.00274378} = \sqrt{0.02915017} = 0.1707342$. Thus, the 95% confidence interval for β_1 is

$$1.61591 \pm 2.074 \cdot 0.1707342 = 1.61591 \pm 0.3541027 = (1.261807, 1.970013)$$

In conclusion, for β_2 , we have the following standard error $\sqrt{S^2 \cdot 0.000001229} = 0.003613448$ and thus the 95% confidence interval for β_2 is

$$0.01438 \pm 2.074 \cdot 0.003613448 = 0.01438 \pm 0.007494291 = (0.006885709, 0.02187429)$$

As one can see in every confidence interval, the zero is not included, thus all the coefficients are statistically significant at 5%.