## MTH5126 - Statistics for Insurance

## Worksheet 3 - Solutions

## Q1. Compound distribution

$S$ has a compound distribution with Poisson parameter 4. The individual claim amounts are either 1 , with probability 0.3 , or 3 , with probability 0.7 .
Calculate the probability that $S=4$.
We need to consider how we could get to an aggregate claim amount of 4 .
This could happen in two ways:

2 claims, one for 1 and one for 3 .
4 claims, all for an amount of 1 .

The probability of this happening is therefore:

$$
\begin{array}{r}
P(S=4)=P(N=2) P\left(X_{1}=1\right) P\left(X_{2}=3\right)+P(N=2) P\left(X_{1}=3\right) P\left(X_{2}=1\right) \\
+P(N=4) P\left(X_{1}=1\right) P\left(X_{2}=1\right) P\left(X_{3}=1\right) P\left(X_{4}=1\right)
\end{array}
$$

Since the $X_{i}$ 's are identical this simplifies to:

$$
\begin{aligned}
P(S=4) & =2 P(N=2) P(X=1) P(X=3)+P(N=4)[P(X=1)]^{4} \\
& =2 \times \frac{e^{-4} 4^{2}}{2!} \times 0.3 \times 0.7+\frac{e^{-4} 4^{4}}{4!} \times 0.3^{4}=0.06312
\end{aligned}
$$

## Q2. Moments of compound distributions

An insurance portfolio contains policies for three categories of policyholder: $\mathrm{A}, \mathrm{B}$ and C . The number of claims in a year, $N$, on an individual policy follows a Poisson distribution with mean $\lambda$. Individual claim sizes are assumed to be exponentially distributed with mean 4 and are independent from claim to claim. The distribution of $\lambda$, depending on the category of the policyholder, is:

| Category | Value of $\lambda$ | Proportion of policyholders |
| :--- | :--- | :--- |
| A | 2 | $20 \%$ |
| B | 3 | $60 \%$ |
| C | 4 | $20 \%$ |

Denote by S the total amount claimed by a policyholder in oneyear.

1. Prove that $E(S)=E[E(S \mid \lambda)]$
2. $\quad$ Show that $E(S \mid \lambda)=4 \lambda$ and $\operatorname{Var}(S \mid \lambda)=32 \lambda$
3. Calculate $E(S)$
4. Calculate $\operatorname{Var}(S)$
5. Let $f(s)$ denote the marginal probability density for $S$ and let $f(s \mid \lambda)$ denote the conditional probability density for $S \mid \lambda$.

Starting with the RHS of the equation:

$$
\begin{aligned}
& E[E(S \mid \lambda)] \\
& =E\left[\int_{0}^{\infty} s f(s \mid \lambda) d s\right] \\
& =\sum_{i=1}^{3} p\left(\lambda_{i}\right) \int_{0}^{\infty} s f\left(s \mid \lambda_{i}\right) d s \\
& =\int_{0}^{\infty} s \sum_{i=1}^{3} p\left(\lambda_{i}\right) f\left(s \mid \lambda_{i}\right) d s, \text { switching integration and summation }
\end{aligned}
$$

But $\sum_{i=1}^{3} p\left(\lambda_{i}\right) f\left(s \mid \lambda_{i}\right)=f(s)$ by definition. and so:

$$
E(E(S \mid \lambda))=\int_{0}^{\infty} s f(s) d s=E(S)
$$

2. Using the results for compound distributions we get:
$E(S \mid \lambda)=E(N \mid \lambda) E(X \mid \lambda)$, using formula for the mean of compound distributions

$$
\begin{aligned}
& =E(N \mid \lambda) E(X), \text { since } X \text { is independent of } \lambda \\
& =\lambda .4=4 \lambda
\end{aligned}
$$

$\operatorname{Var}(S \mid \lambda)=E(N \mid \lambda) \operatorname{Var}(X \mid \lambda)+\operatorname{Var}(N \mid \lambda)[E(X \mid \lambda)]^{2}$, using formula for the variance of compound distributions

$$
\begin{aligned}
& =E(N \mid \lambda) \operatorname{Var}(X)+\operatorname{Var}(N \mid \lambda)[E(X)]^{2}, \text { since } X \text { is independent of } \lambda \\
& =\lambda \times 16+\lambda \times 4^{2} \\
& =32 \lambda
\end{aligned}
$$

3. 

$E(S)=E[E(S \mid \lambda)]$, using results from part 1(by the law of total expectation)
$=E(4 \lambda)$, using results from part 2
$=4 E(\lambda)$
$=4 \times(0.2 \times 2+0.6 \times 3+0.2 \times 4)$
$=4 \times 3$
$=12$
4. First note that $E(\lambda)=3$ and
$\operatorname{Var}(\lambda)=E\left(\lambda^{2}\right)-[E(\lambda)]^{2}=0.2 \times 2^{2}+0.6 \times 3^{2}+0.2 \times 4^{2}-3^{2}=0.4$
$\operatorname{Var}(S)=\operatorname{Var}[E(S \mid \lambda)]+E[\operatorname{Var}(S \mid \lambda)]$, by the law of total variance
$=\operatorname{Var}(4 \lambda)+E(32 \lambda)$, using results from part 2
$=16 \times \operatorname{Var}(\lambda)+32 E(\lambda)$
$=16 \times 0.4+32 \times 3$
$=102.4$

## Q3. R

Before answering this question, generate the vector, $X$, in R using the following code:

```
set.seed(1027); X = rexp(n=1000, rate=0.01)
```

The vector $X$ represents the gross claim sizes of 1,000 claims. The payments are to be split between an insurance company and its reinsurer under an Excess of Loss reinsurance arrangement with a retention level $M=400$.
(i) Determine the proportion of the claims that are fully covered by the insurer.

Hint: The following code might help.

```
length(X[X<=M])/length(X)
```

(ii) Generate an additional vector, $Y$, which is of the same length as $X$, such that $Y$ represents the amounts to be paid by the insurer for each component of $X$.

Hint: Use the pmin function.
(iii) Generate an additional vector, $Z$, which is of the same length as $X$, such that $Z$ represents the amounts to be paid by the reinsurer for each component of $X$.

An actuary assumes that the underlying gross claims distribution follows an exponential distribution of some unknown rate $\lambda$. The actuary needs to estimate $\lambda$ using only the claim amounts recorded in vector $Y$.
(iv) Construct R code that calculates the log-likelihood, as a function of the parameter $\lambda$, given the claim amounts data in vector $Y$.

Hint: This is estimation when sample is censored, see lecture slides.
(v) Using the function nlm , determine the value of $\lambda$ at which the log-likelihood function reaches its maximum.

Hint: The nlm function performs minimisation, not maximisation. However, maximising the log-likelihood function is the same as minimising the negative loglikelihood. So, we first define the function that we want to hand to nlm to be minimised.

Solution:

```
#Q(i) Proportion of claims fully covered by the insurer
set.seed(1027)
X=rexp (1000,0.01)
M=400
> length(X[X<=M])/ length(X)
[1] 0.987
```

So the proportion of claims fully covered by the insurer is $98.7 \%$.

```
#Q(ii) Vector Y, same length as X, represents the amounts to
be paid by the insurer for each component of X.
```

```
Y=pmin(X,M)
```

The following code and output show that $Y$ is indeed the same length as $X$, i.e. the length of Y is also 1000 .

```
> length(Y)
```

[1] 1000

```
#Q(iii) Vector Z represents the amounts to be paid by the
reinsurer for each component of X.
```

$Z=X-Y$

The following code and output show that Z is indeed the same length as X , i.e. the length of Z is also 1000 .

```
> length(Z)
```

[1] 1000

```
#OR
```

$Z=\operatorname{pmax}(0, X-M)$

The following code and output show that Z is indeed the same length as X , i.e. the length of Z is also 1000 .

```
> length(Z)
[1] 1000
#Q(iv) Sample is censored. See lecture notes on how the
complete likelihood function is made up of two parts.
#The first part relates to the }987\mathrm{ claims, the second part
relates to the 13 claims.
#We assume all claims are independent.
```

\#So the likelihood function for 987 claims $=$
(lambda^987)*exp (-lambda*sum_of_987_claims)
\#And the likelihood function for 13 claims $=[P(X>M)]^{\wedge} 13=$ [exp (-lambda*M*13)
\#And the complete likelihood function, L = Product of the two likelihood functions above
\#Note that sum_of_987_claims $+13 * M$ is simply the sum of all the components in vector $Y$.

Final answer is:
$S=\operatorname{sum}(Y)$
logLikelihood<-function(lambda) \{
987*log(lambda) -lambda*S
\#Q(v) Find the value of lambda at which the logLikelihood is at its maximum.
\#The following graph plotting is not required by the question but it helps us to think about where the maximum is.
\#Plot the logLikelihood just to get an idea of how it looks like.
\#If necessary, adjust the graph so that we can roughly see where the max is.
\#We see this happens at around lambda $=0.01$.
1 ambda $=\operatorname{seq}(0.0001,0.05, b y=0.0001)$
plot(lambda, logLikelihood(lambda))

\#We find lambda using numerical algorithm such as nlm. \#Note that $n l m$ performs minimisation, not maximisation. \#However, maximising logLikelihood is the same as minimising the -logLikelihood.
\#So we define the function that we want to hand to nlm to be minimised.

Function $=$ function(lambda) \{
-logLikelihood(lambda)
\}
\#To find out more about nlm, we run the following and look at the notes under Help.
? nlm
\#p is our starting value for the iterative algorithm. From the graph we know the max is around 0.01 , so set $p=0.01$.
nlm(f=Function, $\mathrm{p}=0.01$ )
$>$ nlm(f=Function, $\mathrm{p}=0.01$ ) \$estimate
[1] 0.01023209
So, the estimate for lambda is 0.01023209 .

