

# Topic : UNIT ROOT NONSTATIONARITY

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CONSIDER AR(1) MODEL :

$$X_t = \phi X_{t-1} + \varepsilon_t$$

- IF  $|\phi| < 1$  then THE MODEL HAS A STATIONARY SOLUTION
- IF  $\phi = 1$  then THE MODEL IS A UNIT ROOT MODEL WHICH IS NON-STATIONARY



SUCH MODEL DOES NOT SATISFY  
CONDITION OF COVARIANCE STATIONARITY  
OF AN AR(1) MODEL

A time series  $X_t$  is covariance stationary if :

$$E(X_t) = \mu_x$$

$$\text{Var}(X_t) = \sigma_x^2$$

$$\text{Cov}(X_t, X_{t-k}) = \gamma_k$$

IN THIS CLASS WE ARE GOING TO CONSIDER THE RANDOM WALK AND THE RANDOM WALK WITH DRIFT MODELS.

- A TIME SERIES  $X_t$  IS CALLED A RANDOM WALK IF  $X_t = X_{t-1} + \epsilon_t$   $t=0,1,\dots,$  where  $\epsilon_t$  IS A ZERO MEAN WHITE NOISE

- A RANDOM WALK WITH A DRIFT CAN BE WRITTEN AS:  $X_t = \mu + X_{t-1} + \epsilon_t$

→ 'A RANDOM-WALK IS COVARIANCE NON-STATIONARY MODEL

## Problem 7.1

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$$X_t = \mu + X_{t-1} + \varepsilon_t \quad t=1, 2, \dots$$

where  $X_0 = 0$

$\varepsilon_t$  IS WN  $(0, \sigma_\varepsilon^2)$

$X_t$  CAN BE REWRITTEN AS:

$$X_t = \mu + \underbrace{X_{t-1}}_{\downarrow} + \varepsilon_t$$

$$= \mu + \underbrace{(\mu + X_{t-2} + \varepsilon_{t-1})}_{\downarrow} + \varepsilon_t$$

$$= 2\mu + \underbrace{X_{t-2}}_{\downarrow} + \varepsilon_{t-1} + \varepsilon_t$$

$$= 2\mu + \underbrace{(\mu + X_{t-3} + \varepsilon_{t-2})}_{\downarrow} + \varepsilon_{t-1} + \varepsilon_t$$

$$= 3\mu + X_{t-3} + \varepsilon_{t-2} + \varepsilon_{t-1} + \varepsilon_t$$

$$= \dots$$

$$= t\mu + \underbrace{X_0}_{=0} + \varepsilon_t + \varepsilon_{t-1} + \varepsilon_{t-2} + \dots + \varepsilon_1$$

we assumed

$$X_0 = 0$$

$\varepsilon_1$  IS  
COMING  
FROM  
 $\varepsilon_{t-(t-1)}$

$$= t\mu + \varepsilon_t + \varepsilon_{t-1} + \varepsilon_{t-2} + \dots + \varepsilon_1$$

(i)  $E(X_t) = t\mu$

$$E(X_t) = E(t\mu + \epsilon_t + \epsilon_{t-1} + \epsilon_{t-2} + \dots + \epsilon_1)$$

$$= \underbrace{E(t\mu)}_{=t\mu} + \underbrace{E(\epsilon_t)}_{=0} + \underbrace{E(\epsilon_{t-1})}_{=0} + \dots + \underbrace{E(\epsilon_1)}_{=0}$$

$\epsilon_t$  IS A WHITE NOISES SEQUENCE WITH ZERO MEAN

so

$$E(X_t) = t\mu$$

(ii)  $Var(X_t) = t \sigma_\epsilon^2$

$$Var(X_t) = E[(X_t - E(X_t))^2] =$$

$$= E\left[\left(\underbrace{(t\mu + \epsilon_t + \epsilon_{t-1} + \dots + \epsilon_1)}_{\downarrow} - t\mu\right)^2\right]$$

$$= E\left[(\epsilon_t + \epsilon_{t-1} + \epsilon_{t-2} + \dots + \epsilon_1)^2\right]$$

$$= \underbrace{E[\epsilon_t^2]}_{=\sigma_\epsilon^2} + \underbrace{E[\epsilon_{t-1}^2]}_{=\sigma_\epsilon^2} + \dots + \underbrace{E[\epsilon_1^2]}_{=\sigma_\epsilon^2}$$

$$Var(\epsilon_t) = E\left[(\epsilon_t - \underbrace{E(\epsilon_t)}_{=0})^2\right] = E[\epsilon_t^2]$$

$$= \sigma_\epsilon^2 + \sigma_\epsilon^2 + \dots + \sigma_\epsilon^2$$

$$= t \sigma_\epsilon^2$$

(iii)  $Cov(X_t, X_s) = \min(t, s) \sigma_\epsilon^2$

WE KNOW  $X_t = \mu + X_{t-1} + \epsilon_t$

$X_t = t\mu + \epsilon_t + \epsilon_{t-1} + \epsilon_{t-2} + \dots + \epsilon_1$

AND  $E(X_t) = t\mu$

WE CAN NOW DEFINE  $X_s$  AS

$X_s = s\mu + \epsilon_s + \epsilon_{s-1} + \dots + \epsilon_1$

where  $t \geq s$

WE OBTAIN THAT  $E(X_s) = s\mu$

• BY THE DEFINITION OF COVARIANCE:

$$\begin{aligned}
Cov(X_t, X_s) &= E\left[ \underbrace{(X_t - E(X_t))}_{\text{red}} \underbrace{(X_s - E(X_s))}_{\text{purple}} \right] = \\
&= E\left[ \left( (t\mu + \epsilon_t + \epsilon_{t-1} + \dots + \epsilon_1) - t\mu \right) \cdot \right. \\
&\quad \left. \cdot \left( (s\mu + \epsilon_s + \epsilon_{s-1} + \dots + \epsilon_1) - s\mu \right) \right] = \\
&= E\left[ \underbrace{(\epsilon_t + \epsilon_{t-1} + \dots + \epsilon_1)}_{\text{red}} (\epsilon_s + \epsilon_{s-1} + \dots + \epsilon_1) \right] =
\end{aligned}$$

IT CAN BE WRITTEN AS:

$(\{\epsilon_s + \epsilon_{s-1} + \dots + \epsilon_1\} + \{\epsilon_{s+1} + \epsilon_{s+2} + \dots + \epsilon_t\})$

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$$= E \left[ \left( \{ \varepsilon_s + \varepsilon_{s-1} + \dots + \varepsilon_1 \} + \{ \varepsilon_{s+1} + \dots + \varepsilon_t \} \right) \cdot \left( \varepsilon_s + \varepsilon_{s-1} + \dots + \varepsilon_1 \right) \right] =$$

$$= E \left[ \left( \varepsilon_s + \varepsilon_{s-1} + \dots + \varepsilon_1 \right)^2 + \left( \varepsilon_{s+1} + \dots + \varepsilon_t \right) \left( \varepsilon_s + \varepsilon_{s-1} + \dots + \varepsilon_1 \right) \right] =$$

$$= \underbrace{E \left[ \left( \varepsilon_s + \varepsilon_{s-1} + \dots + \varepsilon_1 \right)^2 \right]}_{= 5\sigma_\varepsilon^2} + \underbrace{E \left[ \left( \varepsilon_{s+1} + \dots + \varepsilon_t \right) \left( \varepsilon_s + \varepsilon_{s-1} + \dots + \varepsilon_1 \right) \right]}_{= 0} =$$

BECAUSE

$$\begin{aligned} & E(\varepsilon_s^2) + E(\varepsilon_{s-1}^2) + \dots + E(\varepsilon_1^2) = \\ & = \sigma_\varepsilon^2 + \sigma_\varepsilon^2 + \dots + \sigma_\varepsilon^2 = \\ & = 5\sigma_\varepsilon^2 \end{aligned}$$

BECAUSE

$$\begin{aligned} & E \left[ \varepsilon_{s+1} \varepsilon_s + \varepsilon_{s+1} \varepsilon_{s-1} + \dots + \varepsilon_t \varepsilon_1 \right] = \\ & = E \left[ \varepsilon_{s+1} \varepsilon_s \right] + E \left[ \varepsilon_{s+1} \varepsilon_{s-1} \right] + \dots = \\ & = 0 + 0 + \dots = 0 \end{aligned}$$

$\varepsilon$  IS A WHITE NOISE SEQUENCE  
AND THEREFORE

$$E(\varepsilon_i \varepsilon_j) = 0 \quad \text{if } i \neq j$$

$$= 5\sigma_\varepsilon^2 + 0 =$$

$$= 5\sigma_\varepsilon^2$$

WE CAN CONCLUDE :

$$\text{COV}(X_t, X_s) = 5\sigma_\varepsilon^2 = \min(t, s) \sigma_\varepsilon^2$$

Additional Notes:

$\epsilon_t + \epsilon_{t-1} + \epsilon_{t-2} + \dots + \epsilon_1$  CAN BE WRITTEN AS

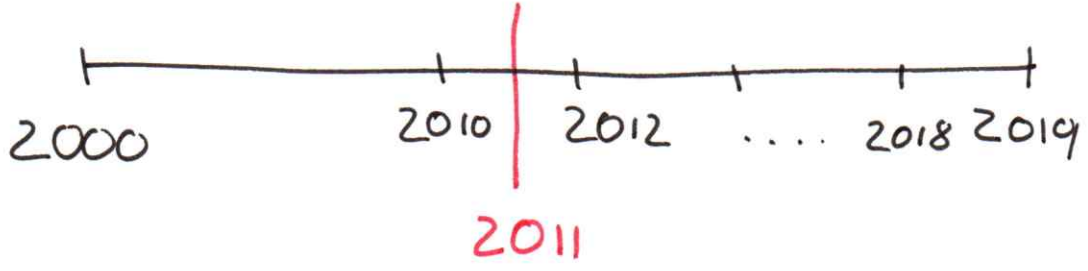
$$\{\epsilon_s + \epsilon_{s-1} + \dots + \epsilon_1\} + \{\epsilon_{s+1} + \epsilon_{s+2} + \dots + \epsilon_{t-1} + \epsilon_t\}$$

WE CAN EXPLAIN IT USING AN EXAMPLE:

WE CONSIDER THE TIME PERIOD 2000 - 2019



AND WE WANT TO TEST THE CORRELATION BETWEEN  $X_{2019}$  AND  $X_{2011}$



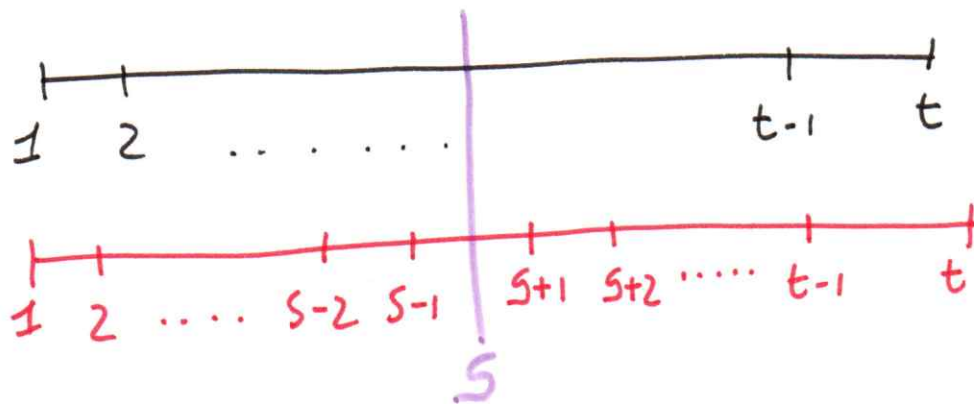
the sample could be split in two

WE CAN GENERALIZED IT BY USING  $t$

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AND WE WANT TO COMPUTE THE CORRELATION BETWEEN  $X_t$  AND  $X_s$  where  $t > s$



SO THE PERIOD  $1 - t$  CAN BE DIVIDED INTO TWO PERIODS:

$$(1 - s) \text{ AND } (s+1 - t)$$

THE ABOVE EXAMPLE EXPLAINS WHY

$$\begin{aligned} \epsilon_t + \epsilon_{t-1} + \epsilon_{t-2} + \dots + \epsilon_1 &= \\ &= (\epsilon_1 + \dots + \epsilon_{s-1} + \epsilon_s) + (\epsilon_{s+1} + \dots + \epsilon_{t-1} + \epsilon_t) \end{aligned}$$



## Problem 7.2

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CONSIDER  $X_t = X_{t-1} + \varepsilon_t$

WITH INITIAL VALUE  $X_0 = 0$  AND

$\varepsilon_t$  AS A WHITE NOISE WITH  
ZERO MEAN AND VARIANCE  $\sigma_\varepsilon^2$

WE KNOW  $X_t = X_{t-1} + \varepsilon_t$

THEN  $X_{t+1} = X_t + \varepsilon_{t+1}$

AND  $X_{t+2} = X_{t+1} + \varepsilon_{t+2}$

(i) OBTAIN THE 1-STEP AHEAD PREDICTION  $\hat{X}_t(1)$

THE 1-STEP AHEAD FORECAST IS DEFINED BY:

$$\hat{X}_t(1) = E[X_{t+1} | F_t] = [X_{t+1}]$$

WE HAVE THAT:

$$\hat{X}_t(1) = E[X_{t+1} | F_t] = [X_{t+1}] =$$

$$= [X_t + \varepsilon_{t+1}] =$$

$$= \underbrace{[X_t]}_{= X_t} + \underbrace{[\varepsilon_{t+1}]}_{= 0} = X_t$$

BECAUSE  
 $X_t$  IS KNOWN  
WHEN WE KNOW  $F_t$

BECAUSE  $\varepsilon_{t+1}$  IS INDEPENDENT  
OF THE "HISTORY"  $F_t$

(ii) DERIVE 2-STEP AHEAD PREDICTION  $\hat{X}_t(2)$

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$$\begin{aligned}\hat{X}_t(2) &= E[X_{t+2} | F_t] = [X_{t+2}] = \\ &= [X_{t+1} + \varepsilon_{t+2}] = \\ &= \underbrace{[X_{t+1}]}_{= \hat{X}_t(1)} + \underbrace{[\varepsilon_{t+2}]}_{= 0} \\ &= X_t\end{aligned}$$

so  $\hat{X}_t(2) = X_t + 0 = X_t$

THE FORECAST OF A RANDOM-WALK MODEL IS THE LAST OBSERVATION  $X_t$

FOR ANY K-STEP AHEAD FORECAST

$$\hat{X}_t(k) = X_t$$

THEREFORE :

FORECASTING IS NOT MEAN REVERTING

(iii) IS THE FORECAST MEAN REVERTING?

NO, THE FORECAST IS NOT MEAN REVERTING

BECAUSE  $E(X_t) = 0$  WHICH IS DIFFERENT FROM  $X_t$

WE KNOW  $X_t = X_{t-1} + \epsilon_t$

$$X_t = (X_{t-2} + \epsilon_{t-1}) + \epsilon_t$$

$$X_t = X_{t-3} + \epsilon_{t-2} + \epsilon_{t-1} + \epsilon_t$$

⋮

$$X_t = \underbrace{X_0}_{=0} + \epsilon_1 + \dots + \epsilon_{t-1} + \epsilon_t$$

BECAUSE INITIAL VALUE  $X_0 = 0$

$$\begin{aligned} E(X_t) &= E(\epsilon_1 + \epsilon_2 + \dots + \epsilon_t) \\ &= E(\epsilon_1) + \dots + E(\epsilon_t) = \\ &= 0 \end{aligned}$$

THIS IS DIFFERENT FROM:

$$\hat{X}_t(1) = X_t$$

$$\hat{X}_t(2) = X_t$$

⋮

$$\hat{X}_t(K) = X_t$$

# Problem 7.3

CONSIDER  $X_t = X_{t-1} + Z_t$

where  $Z_t$  IS AN AR(2) MODEL:

$$Z_t = 3 + 0.5 Z_{t-1} + 0.1 Z_{t-2} + \epsilon_t$$

where  $\epsilon_t$  IS WN  $(0, \sigma_\epsilon^2)$

• FIND 1-STEP AHEAD FORECAST  $\hat{X}_t(1)$

• FIRST, WE NEED TO FIND 1-STEP AHEAD FORECAST OF  $Z_{t+1}$

$$\hat{Z}_t(1) = E[Z_{t+1} | F_t] = [Z_{t+1}]$$

$$= [3 + 0.5 Z_t + 0.1 Z_{t-1} + \epsilon_{t+1}]$$

$$= 3 + 0.5 [Z_t] + 0.1 [Z_{t-1}] + [\epsilon_{t+1}] =$$

$Z_t$  AND  $Z_{t-1}$   
ARE KNOWN WHEN  
WE KNOW  $F_t$

$Z_t$   
.....  
 $Z_{t-1}$

BECAUSE  $\epsilon_{t+1}$  IS  
INDEPENDENT OF  
THE HISTORY OF  $F_t$

$$= 3 + 0.5 Z_t + 0.1 Z_{t-1} =$$

$$= 3 + 0.5 (X_t - X_{t-1}) + 0.1 (X_{t-1} - X_{t-2})$$

WE KNOW  $X_t = X_{t-1} + Z_t$  THEN  $Z_t = X_t - X_{t-1}$

• THEN

$$\hat{X}_t(1) = E[X_{t+1} | F_t] = [X_{t+1}]$$

$$= [X_t + Z_{t+1}] =$$

$$= \underbrace{[X_t]}_{= X_t} + \underbrace{[Z_{t+1}]}_{= \hat{Z}_t(1)} =$$

$$= \hat{Z}_t(1)$$

$$= 3 + 0.5(X_t - X_{t-1}) + 0.1(X_{t-1} - X_{t-2})$$

$$= X_t + 3 + 0.5(X_t - X_{t-1}) + 0.1(X_{t-1} - X_{t-2}) =$$

$$= \underline{X_t} + 3 + \underline{0.5 X_t} - \underline{0.5 X_{t-1}} + \underline{0.1 X_{t-1}} - 0.1 X_{t-2} =$$

$$= 3 + 1.5 X_t - 0.4 X_{t-1} - 0.1 X_{t-2}$$

## Problem 7.4

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STEPS TO FOLLOW:

- PLOT GDP SERIES
- GENERATE LOG(GDP) AND PLOT THE NEW SERIES
- TEST FOR UNIT ROOT

$H_0$ : SERIES CONTAINS UNIT ROOT

$H_1$ : SERIES IS STATIONARY

SINCE P-VALUE  $> \alpha$  THEN  
WE FAIL TO REJECT  $H_0$

→ LOGGDP HAS AT LEAST A UNIT ROOT

- OUTPUT SHOWS WE CAN FIT THE MODEL:

$$\Delta Y_t = c + \text{"trend"} t + \beta_c Y_{t-1} + \delta_1 \Delta Y_{t-1} + \delta_2 \Delta Y_{t-2} + \epsilon_t$$

where  $\Delta Y_t = Y_t - Y_{t-1}$

AND  $Y_t$  IS LOGGDP

OUTPUT SUGGESTS THAT:

$$\Delta Y_t = 0.4221 \Delta Y_{t-1} + 0.1941 \Delta Y_{t-2} + \epsilon_t$$

THE COEFFICIENTS  $c$ , trend,  $\beta_c$ ,  $\delta_3$ ,  $\delta_4$   
ARE NOT SIGNIFICANT

- THE SERIES IS NOT STATIONARY SO  
WE TEST THE FIRST DIFFERENCE

$$\text{LOGGDP}_t - \text{LOGGDP}_{t-1}$$

THE FIRST DIFFERENCE IS STATIONARY

→ P-VALUE  $< \alpha$  THEN WE REJECT  $H_0$