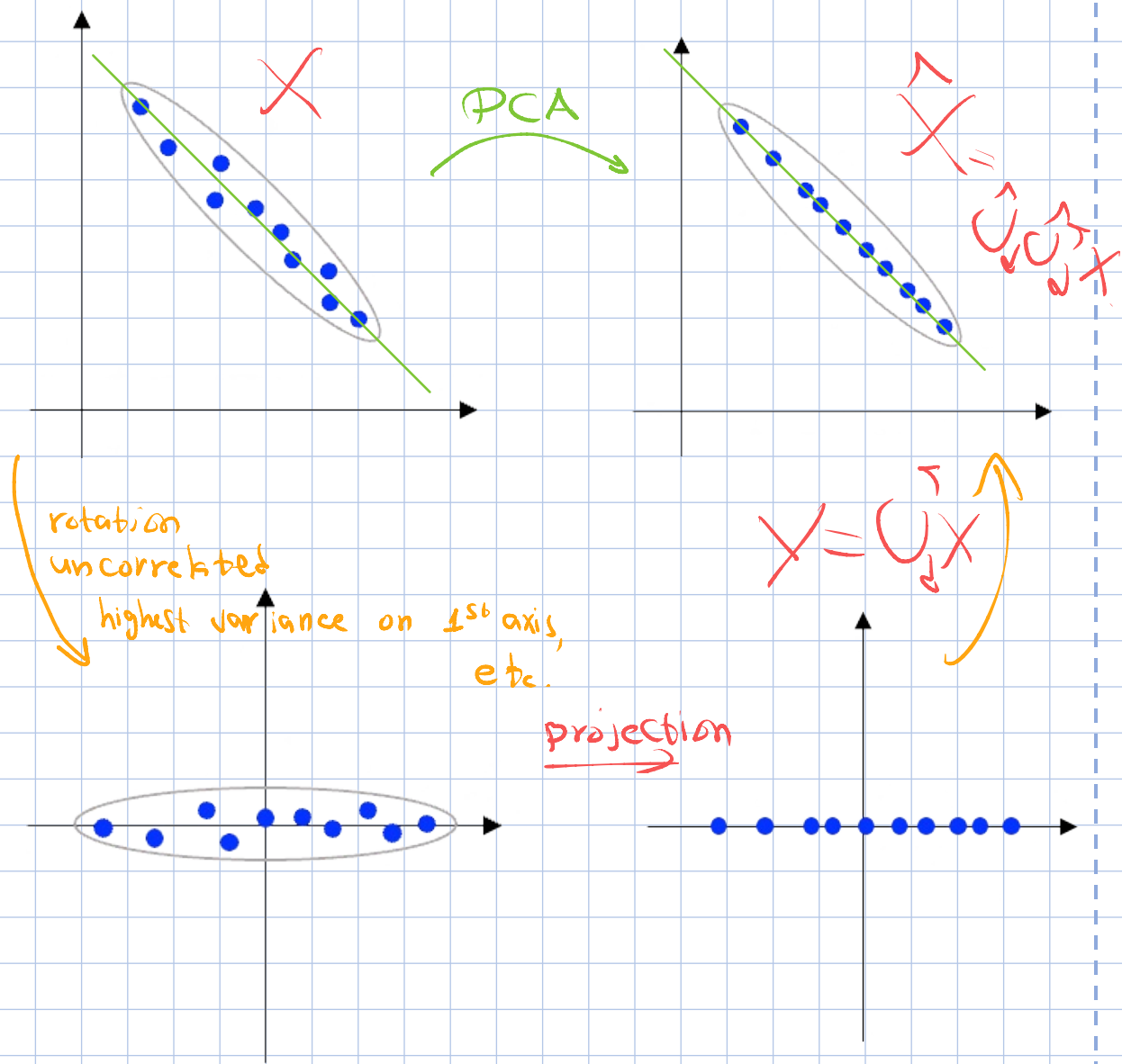


# PRINCIPAL COMPONENT ANALYSIS (PCA)



INPUT:  $x_1, x_2, \dots, x_n \in \mathbb{R}^D$

GOAL: Fit a lower-dimensional  $V$  space of dim  $d \ll D$  to  $x_1, \dots, x_n$  affine.

TWO VIEWS:

- (1) statistical
- (2) geometric.

## THE STATISTICAL VIEW

Suppose:  $\underline{x} = (x^{(1)}, x^{(2)}, \dots, x^{(D)})$  - random vector in  $\mathbb{R}^D$

• assume:  $E(\underline{x}) = \underline{0}$

•  $\Lambda_{\underline{x}} = \text{Cov}(\underline{x}) \Rightarrow (\Lambda_{\underline{x}})_{i,j} = \text{Cov}(x^{(i)}, x^{(j)})$

covariance matrix.

GOAL: Find a projection from  $\underline{x} \in \mathbb{R}^D$  to  $\underline{y} \in \mathbb{R}^d$  ( $d < D$ ) such that:

$1 \leq i \leq d$

$$(1) \quad y^{(i)} = \underline{u}_i^T \underline{x} \quad \underline{u}_i \in \mathbb{R}^D, \quad \underline{u}_i^T \underline{u}_i = 1$$

*non-random*

(2)  $y^{(1)}, y^{(2)}, \dots, y^{(d)}$  - uncorrelated

(3)  $y^{(1)}, \dots, y^{(d)}$  - maximum possible variance

(4)  $\text{Var}(y^{(1)}) \geq \text{Var}(y^{(2)}) \geq \dots \geq \text{Var}(y^{(d)})$

Find  $y^{(1)}$  ( $\underline{u}_1$ ):

$$\underline{u}_1 = \underset{\underline{u}}{\operatorname{argmax}} \operatorname{Var}(\underline{u}^T \cdot \underline{x}) \quad \|\underline{u}\|=1$$

$(\underline{u}^T \underline{u} = 1)$

$$\otimes \operatorname{Var}(\underbrace{\underline{u}^T \underline{x}}_{\text{scalar}}) = E((\underline{u}^T \underline{x})^2) =$$

$$= E((\underline{u}^T \underline{x})(\underline{u}^T \underline{x})^T)$$

$$= E(\underline{u}^T (\underline{x} \underline{x}^T) \underline{u})$$

$$= \underline{u}^T E(\underline{x} \underline{x}^T) \cdot \underline{u}$$

$$= \underline{u}^T \Delta_x \underline{u}$$

$$\Rightarrow \underline{u}_1 = \underset{\underline{u}}{\operatorname{argmax}} \underline{u}^T \Delta_x \underline{u}, \quad \|\underline{u}\|=1$$

CLAIM:

Let  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_D$  - eig. vals of  $\Delta_x$

then:  $\underline{u}_1$  = eigen vector for  $\lambda_1$ .

$$(\Delta_x \underline{u}_1 = \lambda_1 \underline{u}_1)$$

Find  $y^{(2)}$  ( $\underline{u}_2$ ):

$$\underline{u}_2 = \underset{\underline{u}}{\operatorname{argmax}} \operatorname{Var}(\underline{u}^T \cdot \underline{x}) \quad \text{s.t. (1) } \|\underline{u}\|=1$$

$$\otimes (2) \operatorname{Cov}(\underline{u}_1^T \underline{x}, \underline{u}^T \underline{x}) = 0$$

$$\otimes \operatorname{Cov}(\underline{u}_1^T \underline{x}, \underline{u}^T \underline{x}) = \underbrace{\underline{u}_1^T \Delta_x \underline{u}}_{(\Delta_x \underline{u}_1)^T} = \lambda_1 \underline{u}_1^T \underline{u}$$

$$\underline{u}_2 = \operatorname{argmax}_{\underline{u}} \operatorname{Var}(\underline{u}^T \cdot \underline{x}) \quad \text{s.t. (i) } \|\underline{u}\|=1$$
$$\underline{u}_2^T \cdot \underline{u} = 0$$

CLAIM:

$\underline{u}_2$  = eigen-vector for  $\lambda_2$

⋮

↓

$\underline{u}_i$  = eigen-vec. for  $\lambda_i$

## THEOREM

The projection of  $\underline{x}$  on  
the first  $d$  P.C.-s is:  
(principal components)

$$y^{(i)} = \underline{u}_i^T \cdot \underline{x} \quad \underline{u}_i = \text{eigen vec of } \Delta \underline{x} \text{ for } \lambda_i$$

↓ decreasing order

$$\underline{y} = \underline{U}_d^T \cdot \underline{x} \quad \underline{U}_d = \left( \begin{array}{c|c|c|c} \underline{u}_1 & \underline{u}_2 & \dots & \underline{u}_d \\ \downarrow & \downarrow & & \downarrow \end{array} \right)$$

## DCA - For A SAMPLE

Before -  $\underline{x} \in \mathbb{R}^D$  - Random vector.

$$\Delta_x = \underline{\text{known}}$$

In reality -  $\Delta_x$  - unknown

BUT we have data:

$$\underline{x}_1, \underline{x}_2, \dots, \underline{x}_n \in \mathbb{R}^D$$

Assume:  $\bar{\underline{x}}_n = \frac{1}{n} \sum \underline{x}_i = 0$  (centered)

## Empirical Covariance:

$$\hat{\Delta}_n = \frac{1}{n} \sum \underbrace{\underline{x}_j}_{D \times 1} \cdot \underbrace{\underline{x}_j^T}_{1 \times D} \in \mathbb{R}^{D \times D}$$

Take - data matrix:

$$X = \begin{pmatrix} \underline{x}_1 & \underline{x}_2 & \dots & \underline{x}_n \\ \downarrow & \downarrow & & \downarrow \end{pmatrix} \Rightarrow \underline{\underline{D \times n}}$$

$$\hat{\Delta}_n = \frac{1}{n} X \cdot X^T \approx \Delta_x$$

The sample P.C. of  $X$ :

$\hat{u}_1, \hat{u}_2, \dots, \hat{u}_D$  - eigen-vectors of  $\hat{\Delta}_n$

for  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_D$

(highest eigen-values)

$$\hat{y}^{(i)} = \hat{u}_i \cdot \underline{x}$$

$$\underline{y} = \hat{U}_D \cdot \underline{x} \quad \hat{U}_D = \begin{pmatrix} \hat{u}_1 & \hat{u}_2 & \dots & \hat{u}_D \\ \downarrow & \downarrow & & \downarrow \end{pmatrix}$$

## ALGORITHM:

Input: Data =  $\underline{x}_1, \dots, \underline{x}_n \in \mathbb{R}^D$

Generate:  $X = \begin{pmatrix} \underline{x}_1 & \dots & \underline{x}_n \end{pmatrix} \in \mathbb{R}^{D \times n}$

Find eigen-vectors for  $XX^T \rightarrow \underline{\hat{u}}_1, \dots, \underline{\hat{u}}_d$

Generate:  $\hat{U}_d = \begin{pmatrix} \underline{\hat{u}}_1 & \dots & \underline{\hat{u}}_d \end{pmatrix}$

Projection:  $\underline{y} = \hat{U}_d \cdot X \in \mathbb{R}^{d \times n}$

## REMARK:

If  $D \gg n \rightarrow \hat{A}_n \in \mathbb{R}^{D \times D}$

$\frac{1}{n} XX^T$

computing  $\underline{\hat{u}}_1, \dots, \underline{\hat{u}}_d$   
can be slow, etc.

Instead: Take  $X \in \mathbb{R}^{D \times n}$  - smaller matrix

Find SVD:  $X = U \cdot \Sigma^T \cdot V^T$

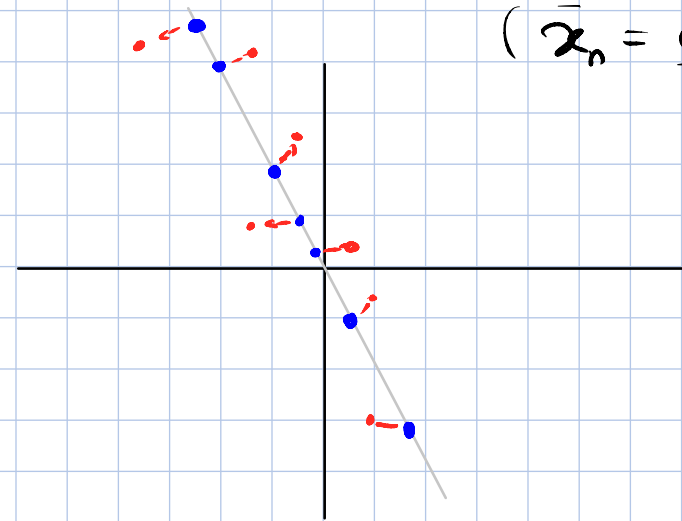
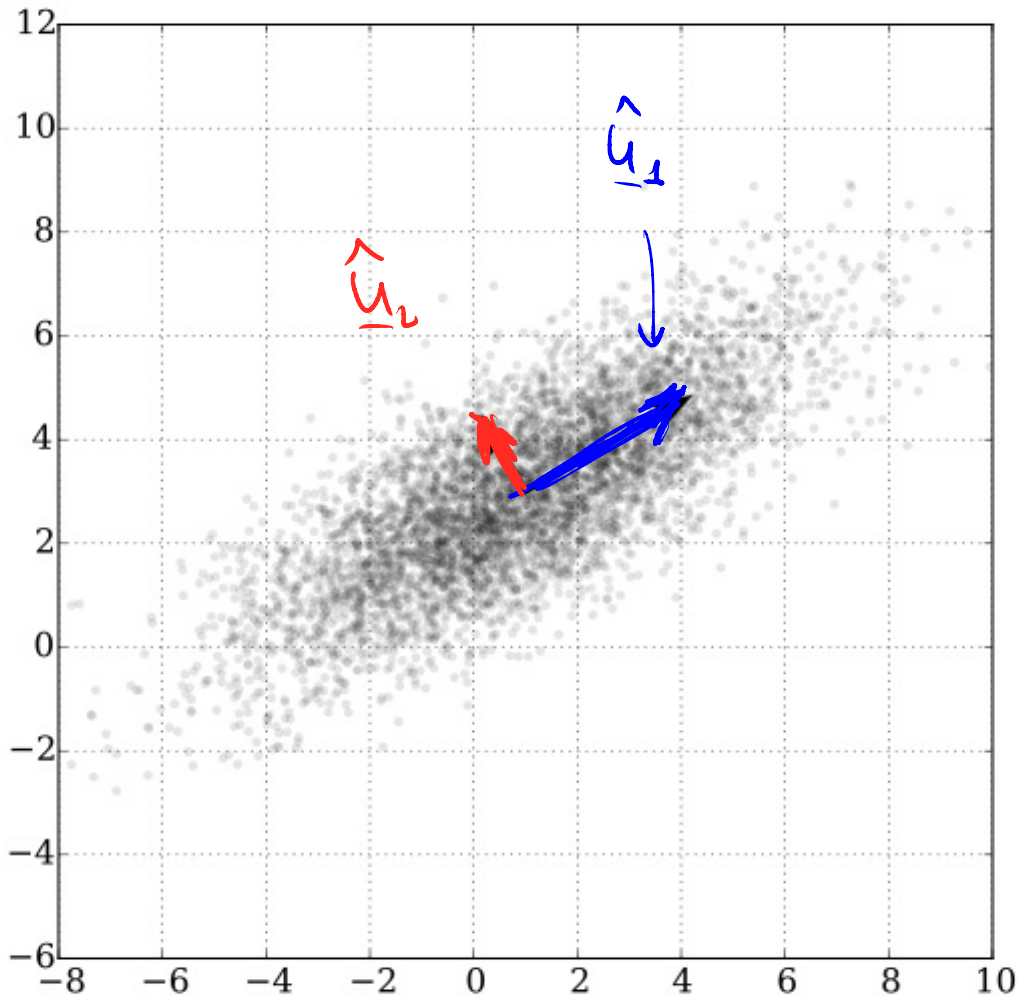
Then the top eigen vectors  
of  $XX^T$   
are the first columns of  $\underline{U}$

## GEOMETRIC VIEW ON PCA


Suppose  $\underline{x}_1, \dots, \underline{x}_n \in \mathbb{R}^D$

all lie on a  $d$ -dimensional  
subspace.

( $\bar{\underline{x}}_n = \underline{0}$ )



$\Rightarrow$  We can find  $U \in \mathbb{R}^{D \times d}$  such that

$$\underline{x}_j = U \cdot \underline{y}_j \quad \underline{y}_j \in \mathbb{R}^d$$


A small diagram showing five blue dots arranged horizontally on a black line, representing the data points in the reduced  $d$ -dimensional space.

In practice:  $\underline{x}_j = U \underline{y}_j + \underline{\varepsilon}_j$   
↘ noise.

Find the "best" U:

⊗  $\arg \min_{\underline{y}_1, \dots, \underline{y}_n} \sum_{j=1}^n \|\underline{x}_j - U \cdot \underline{y}_j\|^2$  s.t.  $U^T U = I$   
 $\sum_{j=1}^n \underline{y}_j = \underline{0}$

THEOREM

The solution for ⊗ is:

- $U = \hat{U}_d = \begin{pmatrix} \hat{u}_1 & \dots & \hat{u}_d \\ \downarrow & & \downarrow \end{pmatrix}$  (top eig-vecs of  $XX^T$ )

- $\hat{\underline{y}}_j = \hat{U}_d^T \cdot \underline{x}_j$

RECONSTRUCTION:

$\underline{x}_j \xrightarrow{\text{project}} \underline{y}_j = \hat{U}_d^T \cdot \underline{x}_j \rightarrow \hat{\underline{x}}_j = \hat{U}_d \cdot \underline{y}_j$   
 $\mathbb{R}^D \quad \mathbb{R}^d \quad \mathbb{R}^D$

$\hookrightarrow \hat{X} = \hat{U}_d \hat{U}_d^T \cdot X$

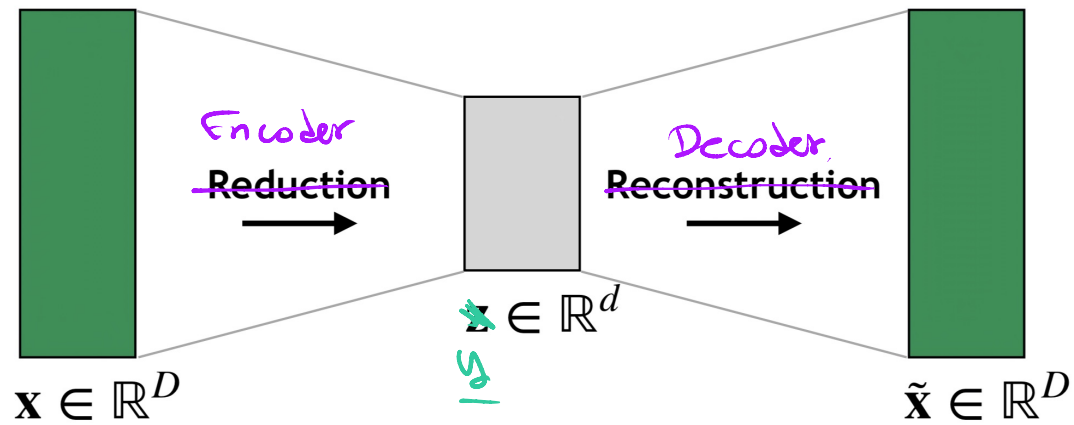
best rank-d approx. for X.

data matrix -  $D \times n$

"  
 $= \sum_{i=1}^d \sigma_i \underline{u}_i \underline{v}_i^T$



## ENCODER - DECODER VIEW



Data:  $X = \begin{pmatrix} x_1 & \dots & x_n \\ \downarrow & & \downarrow \end{pmatrix}$

Encoder:  $Y = E \cdot X$   
matrix

Decoder:  $\tilde{X} = D \cdot Y$

Optimisation problem:

$$\operatorname{argmin}_{D, E} \|X - D \cdot E \cdot X\|_F$$

Solution:

$$D = E^T = \hat{U}_d \text{ - as bef}$$

