

MTH6136 / MTH6136P: Statistical Theory

Duration: 2 hours

Date and time: 31st May 2016, 10:00–12:00

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You should attempt ALL questions. Marks awarded are shown next to the questions.

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Complete all rough workings in the answer book and cross through any work that is not to be assessed.

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Examiner(s): Dr Dudley Stark, Dr Joaquin Miguez

Question 1. Suppose that Y_1, Y_2, \dots, Y_n are independent geometric random variables such that $P(Y = y) = p(1 - p)^{y-1}$ for $y = 1, 2, \dots$. You may assume that the Y_i have expectation $E(Y_i) = 1/p$.

- (a) Find the method of moments estimator for p . [5]
- (b) Show that the distribution of the Y_i belongs to the exponential family of distributions. [5]
- (c) Using (b) and a theorem stated in a lecture, show that $\sum_{i=1}^n Y_i$ is a complete, sufficient statistic for p and that $\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$ is a Minimum Variance Unbiased Estimator for $\phi(p) = 1/p$. [10]

Question 2. Suppose that Y_1, Y_2, \dots, Y_n are independent normal random variables with zero mean and variance θ , where $\theta > 0$.

- (a) Use Neyman's factorisation theorem to show that $\sum_{i=1}^n Y_i^2$ is a sufficient statistic for θ . [6]
- (b) Show that the Cramér-Rao lower bound for unbiased estimators of θ is given by $\text{CRLB}(\theta) = 2\theta^2/n$. [8]
- (c) Prove that $T_n = (Y_2^2 + \dots + Y_{n-1}^2)/(n - 2)$ is an unbiased estimator of θ . Given that $\text{Var}(Y^2) = 2\theta^2$, find the efficiency of T_n and show that the sequence of estimators T_n is asymptotically efficient. [6]

Question 3. Suppose that Y_1, Y_2, \dots, Y_n are independent Weibull random variables with probability density function

$$f_Y(y) = \frac{3y^2}{\theta^3} \exp\left(-\frac{y^3}{\theta^3}\right), \quad y > 0,$$

where $\theta > 0$.

(a) Show that the maximum likelihood estimator of θ is

$$\hat{\theta} = \left(\frac{1}{n} \sum_{i=1}^n Y_i^3\right)^{1/3}.$$

[6]

(b) Given that $E(Y^3) = \theta^3$, show that the Fisher information $E\left(-\frac{\partial^2}{\partial \theta^2} \log L(\theta; Y_1, \dots, Y_n)\right)$ of Y_1, Y_2, \dots, Y_n equals $9n/\theta^2$.

[4]

(c) State a theorem about the limiting distribution (including its parameters) of maximum likelihood estimators as $n \rightarrow \infty$.

[4]

(d) Use your results from parts (b) and (c) to obtain the asymptotic distribution of $\hat{\theta}$, and hence find an approximate $100(1 - \alpha)\%$ confidence interval for θ .

[6]

Question 4. Let Y_1, \dots, Y_n be independent Uniform(0, θ) distributed random variables.

(a) Define what is meant by a *pivot* for θ .

[6]

(b) Let M be the maximum of Y_1, \dots, Y_n . Show that $P(M \leq x) = (x/\theta)^n$ for $0 \leq x \leq \theta$.

[4]

(c) Use (b) to show that M/θ is a pivot for θ .

[4]

(d) Use part (c) to derive an exact 95% confidence interval for θ based on M .

[6]

Question 5. Suppose that Y_1, \dots, Y_n are independent Pascal random variables with probability mass function

$$P(Y = y) = \binom{y+r-1}{r-1} \pi^r (1-\pi)^y, \quad y = 0, 1, \dots,$$

where $0 < \pi < 1$. Consider testing $H_0 : \pi = \pi_0$ against $H_1 : \pi \neq \pi_0$.

- (a) Write down the likelihood, $L(\pi; \underline{y})$, and hence find the generalised likelihood ratio given by $\Lambda(\underline{y}) = \bar{L}(\hat{\pi}_0; \underline{y}) / L(\hat{\pi}; \underline{y})$, where $\hat{\pi}_0$ is the restricted maximum likelihood estimate of π under H_0 and $\hat{\pi}$ is the maximum likelihood estimate. [9]
- (b) State the critical region of the generalised likelihood ratio test in terms of $\Lambda(\underline{y})$ and explain why this only depends on the data through a sufficient statistic. [4]
- (c) Use Wilks' theorem to obtain the critical region of a test with approximate significance level α for large n . [7]

End of Paper.