# Lecture 3A MTH6102: Bayesian Statistical Methods

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Today's lecture will cover

- Review Bayesian updating with continuous parameters and discrete data.
- Construct a posterior for continuous parameters and continuous data.
- Conjugate priors

### **Bayesian** inference

- Suppose we have data y generated from  $p(y | \theta)$  where  $\theta$  is the unknown parameter.
- Start with the prior distribution  $p(\theta)$  about  $\theta$ .
- Likelihood is  $p(y \mid \theta)$ .
- The resulting probability distribution  $p(\theta \mid y)$  is called the posterior distribution.

$$p(\theta \mid y) = \frac{p(\theta) p(y \mid \theta)}{p(y)} \propto p(\theta) p(y \mid \theta)$$

Posterior distribution  $\propto$  prior distribution  $\times$  likelihood

• Our inferences about  $\theta$  are based on this posterior distribution.

- $\theta$  : continuous parameter with prior pdf  $p(\theta)$  and range [a, b].
- x : random discrete data
- likelihood:  $p(x|\theta)$

#### Bayesian updating table

 $p(\theta(x)d\theta =$ 

Hypothesis	prior prob	likelihood	Bayes numerator	posterior prob. $p(\theta x)d\theta$
θ	p( heta)d heta	$p(x \theta)$	p(x  heta)p( heta)d heta	$\frac{p(x \theta)p(\theta)d\theta}{p(x)}$
Total	$\int_a^b p( heta) d heta = 1$		$p(x) = \int_{a}^{b} p(x \theta) p(\theta) d\theta$	1

• The posterior density  $p(\theta|x)$  is obtained by removing  $d\theta$  from the posterior probability in the table. p(x10)p(0/20

- $Y \sim \text{binom}(n, q)$  with unknown binomial probability of success q.
- We observe Y = k successes in *n* trials.  $\mathcal{K} \sim binow (n_1 2)$
- The binomial likelihood p(k|q) for this problem is:

$$P[\underline{l}=\mathcal{L}[2]=p(k \mid q)=\binom{n}{k}q^{k}(1-q)^{n-k}$$

• Convenient prior distribution for q is Beta $(\alpha, \beta)$ :

$$p(q) = rac{q^{lpha - 1}(1 - q)^{eta - 1}}{B(lpha, eta)}$$
 /  $\regin{array}{c}$ 

#### Binomial data/beta prior example

Posterior 
$$\propto$$
 prior  $\times$  likelihood  

$$p(q \mid k) \propto p(q) \times p(k \mid q)$$

$$= \bigoplus_{k=1}^{\infty} \frac{q^{\alpha-1}(1-q)^{\beta-1}}{\beta(\alpha,\beta)} \times \binom{n}{k} q^k (1-q)^{n-k}$$

$$(2)$$

Hence the posterior distribution is proportional to

$$p(q \mid k) \propto q^{k+\alpha-1}(1-q)^{n-k+eta-1}$$

From this, we can recognise that the posterior distribution p(q | k) is the Beta $(k + \alpha, n - k + \beta)$  distribution.

General Beta(a, b) pdf: 
$$f(x) = \frac{x^{a-1}(1-x)^{b-1}}{B(a, b)}$$

We have that the posterior density preseries  

$$p(e|x) = C e^{x+a-1} \cdot (1-e)^{n-x+b-1}$$
  
We see that  $p(e|x)$  is the beta density  
with parameters  $x+a$  and  $n-x+b$   
 $p(e|x| \sim beta (x+a_1n-x+b)$   
The normalising constant  $c$  must be  
 $C = \frac{1}{B(x+a_1n-x+b)}$   
The constant  $c$  makes the density to  
integrate 1

- Data k generated from  $\sim \text{Binom}(n,q)$ , with q unknown
- Continuous hypotheses q in [0, 1].
- Beta $(\alpha, \beta)$  prior p(q)
- Binomial likelihood p(k|q)

Hypothesis	prior prob.	likelihood	Bayes numerator	posterior prob.
q	$Beta(\alpha,\beta)$ dq	binomial(n, q)	$cq^{\scriptscriptstyle k+lpha-1}(1-q)^{\scriptscriptstyle n-k+eta-1}dq$	Beta(k+lpha, n-k+eta)dq
Total	1		$T = \int_0^\infty cq^{k+\alpha-1} (1-q)^{n-k+\beta-1} dq$	1

- The posterior density is  $Beta(k + \alpha, n k + \beta)$
- Note: We don't need to compute T. Once we know the posterior is of the form  $cq^{k+\alpha-1}(1-q)^{n-k+\beta-1}$  we have to find c that makes it a proper density. In this case  $c = 1/\text{Beta}(k + \alpha, n k + \beta)$

## **Conjugate distributions**

· at = initial number of trials



- Binomial likelihood  $k \sim Bin(n, q)$
- Beta $(k + \alpha, n k + \beta)$  posterior distribution for  $q \mid k$
- In this example, we have the same family of distributions for the prior and posterior distribution.
- This is known as a conjugate distribution.
- "The family of Beta distributions is conjugate to the binomial likelihood".

## **Conjugate distributions**

$$\chi \sim \operatorname{Geom}[q]$$
Binomial likelihood:  $p(k \mid q) = \binom{n}{k} q^k (1-q)^{n-k}$ 

$$P[\chi = \chi] = \frac{\chi}{2} (1-q)^{n-k}$$

Beta prior: 
$$p(q) = \frac{q^{\alpha-1}(1-q)^{\beta-1}}{B(\alpha,\beta)}$$

- Considered as functions of q, the prior and likelihood have the same functional form as each other (proportional to  $q^r(1-q)^s$  for some r, s).
- When we multiply them together, we still have the same form.
- This is what characterises conjugate distributions.

0 = probability of heads

• Suppose your prior in the bent coin example is Beta(6,8). You flip the coin 7 times, getting 2 heads and 5 tails. What is the posterior pdf  $p(\theta|x)$ ?



Solution  

$$\alpha = 6_1 \quad \beta = 8$$
  
 $n = 7_1 \quad x = 9_1 \quad n - x = S$   
 $p[0|x] \sim beta(8_113)$ 

- A medical treatment has unknown probability  $\theta$  of success.
- We assume treatment has prior  $f(\theta) \sim \text{Beta}(5,5)$ .
  - Suppose you test it on 10 patients and have 6 successes. Find the posterior distribution on θ. Identify the type of the posterior pdf
     Suppose you recorded the order of the results and got SSSFFSSSFF. Find the posterior based on this new data.

 $D^{6}(1-0)^{4}$ 

Solution  
Prior 
$$p(\theta) = \frac{\theta^{4}(1-\theta)^{4}}{B(515)}$$
,  $\theta \in [01]$   
The data  $x = 6$  successes in 10 patients  
come tram a binomial amodel. So the litelihood  
is just the knomial litelihood  
 $p(6|\theta) = \binom{10}{6} \theta^{6}(7-\theta)^{4}$   
The posterior density,  $p(\theta|6)$  is  
 $p(\theta|6) \propto \frac{\theta^{9}(1-\theta)^{4}}{B(55)} \times \binom{10}{6} \theta^{6}(7-\theta)^{4}$   
 $= C \theta^{10} (1-\theta)^{8}$   
This is the beta  $(11, 9)$ , so the normalising  
constant is just  
 $C = \frac{1}{B(11, 9)}$   
The posterior density is  $p(\theta|6) = \frac{\theta^{10}(1-\theta)^{8}}{B(11, 9)}$ .

(b) The answer is again beta (11,9). The only thing that changes is the binomial coefficient which is just 1. We are now ready to do Bayesian updating when both the parameters and the data take continuous values.

- $\theta$  continuous parameter
- Prior pdf,  $f(\theta)$
- Data: continuous  $x \sim f(x|\theta)$
- Likelihood  $f(x|\theta)$
- posterior pdf,  $f(\theta|\mathbf{x})$
- Bayesian update table

Hypothesis	prior prop	likelihood	Bayes numerator	posterior prop $f(x \theta)d\theta$
$\theta$	$f(\theta)d\theta$	$f(x \theta)$	$f(x \theta) f(\theta) d\theta$	$\frac{f(x \theta)f(\theta)d\theta}{f(x)}$
Total	1		f(x)	1

• 
$$f(x) = \int f(x|\theta) f(\theta) d\theta$$

#### Normal example, known variance

110 •  $y_1, \ldots, y_n \sim N(\mu, \sigma^2).$ 

• It's simpler if only one parameter is unknown.

• First, consider case where only  $\mu$  is unknown. SPOWR

• Is there a conjugate prior for  $\mu$ ?

$$\checkmark \checkmark \checkmark$$

#### Normal example, known variance

- Observed data  $y_1, \ldots, y_n$   $N(\mu, \sigma^2)$  with  $\mu$  unknown and  $\sigma^2$  known.
- Prior distribution  $\mu \sim \mathcal{N}(\mu_0, \sigma_0^2)$ .
- The posterior distribution is

$$\overbrace{\mu \sim \mathcal{N}(\mu_1, \sigma_1^2)}$$

where

$$\mu_{1} = \left(\frac{\mu_{0}}{\sigma_{0}^{2}} + \frac{n\bar{y}}{\sigma^{2}}\right) / \left(\frac{1}{\sigma_{0}^{2}} + \frac{n}{\sigma^{2}}\right)$$
$$\sigma_{1}^{2} = \frac{1}{\sigma_{0}^{2}} + \frac{n}{\sigma^{2}}$$

roof Let  $y_{1, -}, y_n \sim N/\mu_0$  where  $\sigma^2 cs$  known. The lifelihood, plyn, yn lp) ( is the joint density of g1,-, yn: By independence of yn, yn,  $p(y_{1,\cdot},y_{n}|p) = \iint \frac{1}{\sigma \sqrt{2n}} \exp \left\{ \frac{(y_{i}-p_{1})^{2}}{2\sigma^{2}} \right\}$  $= \left(\frac{1}{\sigma \sqrt{2}\sigma}\right)^n \exp \left\{\frac{1}{c=1} - \frac{1}{2\sigma^2} + \frac{(y_c - p_c)^2}{2\sigma^2}\right\}.$  $\propto \exp \xi = \sum_{i=1}^{n} \frac{(y_i - y_i)}{25^2} \xi$ The prior for p is N(po, 50?)  $p(\mu) = \frac{1}{\sqrt{2\pi50^{\circ}}} \exp \frac{5}{250^{\circ}} - \frac{(\mu - \mu_0)^{\circ}}{250^{\circ}}$ we can rewrite the litelihoud as y = 24: N  $p(y|p) \propto e_{\pi p} \sum_{i=1}^{n} \frac{(y_i^2 - 2y_i)(y + p^2)}{2\sigma^2}$  $= exp \left\{ \frac{1}{26a} \left( \frac{Zy}{z} - 2n\overline{y}p + np^{2} \right) \right\}$ 

#### Normal-normal Bayesian update table

• Data: 
$$x \sim \mathcal{N}(\mu, \sigma^2)$$
,  $\sigma^2$  known

- Likelihood:  $f(x|\mu) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2}(x-\mu)^2\}.$
- $\mu$  continuous with prior pdf  $f( heta) \sim \mathcal{N}(\mu_0, \sigma_0^2)$

• posterior 
$$f(\mu|x) \sim \mathcal{N}(\mu_1, \sigma_1^2)$$

Hypothesis	prior prop	likelihood	Bayes numerator	posterior prop $f(x \mu)d\mu$
$\mu$	$\frac{1}{\sqrt{2\pi\sigma_0^2}}\exp\{-\frac{1}{2\sigma_0^2}(\mu-\mu_0)^2\}d\mu$	$\frac{1}{\sqrt{2\pi\sigma^2}}\exp\{-\frac{1}{2\sigma^2}(x-\mu)^2\}$	$c_1 \exp\{-\frac{1}{2\sigma_1^2}(\mu-\mu_1)^2\}d\mu$	$\frac{f(x \mu)f(\mu)d\mu}{f(x)} = \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\{-\frac{1}{2\sigma_1^2}(\mu-\mu_1)^2\}d\mu$
Total	1		$f(x) = \int_{-\infty}^{\infty} c_1 \exp\{-\frac{1}{2\sigma_1^2}(\mu - \mu_1)^2\}d\mu$	1

#### Normal-normal updating formulas

$$a = \frac{1}{\sigma_0^2}, \quad b = \frac{n}{\sigma^2}, \tag{1}$$
$$\mu_1 = \frac{a\mu_0 + b\bar{y}}{a+b}, \quad \sigma_1^2 = \frac{1}{a+b} \tag{2}$$

- The posterior mean  $\mu_1$  is a weighted average of the prior mean  $\mu_0$ and sample average  $\overline{y}$ .
- If *n* is large then the weight *b* is large and  $\bar{y}$  will have a strong influence on the posterior. In fact if  $n \to \infty$ ,  $b/(a+b) \to 1$  and  $a/(a+b) \to 0$ , so  $\mu_1 \to \bar{y}$ .
- If  $\sigma_0^2$  is small then the weight *a* is large and  $\mu_0$  will have a strong influence on the posterior

- Suppose our data follows a  $N(\theta, 1)$  distribution with unknown mean  $\theta$ .
- Suppose our prior on  $\theta$  is N(2,1).
- Suppose we obtain data x = 5
- Compute the Bayesian update table and show that the posterior pdf for  $\theta$  is Normal
- Find the posterior mean and the posterior variance
- Use the updating formulas (1) to find the posterior mean and posterior variance.



prior: blue, posterior: purple, x = 5 (data). The posterior mean lies between the data x = 5 and the prior mean.

### **Board** question



- **(1)** Which plot is the posterior to just the first data value x = 3?
- ② Which plot is the posterior to all 3 data values, x = 3, x = 9 and x = 12?

On a basketball team the free throw percentage over all players is a N(75, 36) distribution. In a given year individual players free throw percentage is  $N(\theta, 16)$  where  $\theta$  is their career average.

This season, Sophie Lee made 85 percent of her free throws.

**(1)** What is the posterior expected values of her career percentage  $\theta$ ?

- The time until failure for a type of light bulb is exponentially distributed with parameter  $\lambda$ .
- We observe *n* bulbs, with failure times  $t = t_1, \ldots, t_n$ .
- The unknown parameter is  $\lambda$ .
- Can we find a conjugate family of distributions for this likelihood?

• A prior is conjugate to a likelihood if the posterior is the same type of distribution as the prior.

	hypothesis	data	prior	likelihood	posterior
Bernoulli/Beta	$ heta \in [0,1]$	X	Beta(lpha,eta)	$Bernoulli(\theta)$	${\sf Beta}(lpha+1,eta)$ or ${\sf Beta}(lpha,eta+1)$
	θ	x = 1	$c_1  heta^{lpha-1} (1- heta)^{b-1}$	θ	$c_3 heta^lpha(1- heta)^{eta-1}$
	θ	x = 0	$c_1  heta^{lpha-1} (1- heta)^{b-1}$	1-θ	$c_3 heta^{lpha-1}(1- heta)^eta$
Binomial/Beta	$ heta \in [0,1]$	X	Beta(lpha,eta)	binomial( $n, \theta$ )	$beta(\alpha + x, \beta + n - x)$
(fixed n)	θ	X	$c_1  heta^{lpha-1} (1- heta)^{b-1}$	$c_2 heta^{x}(1- heta)^{n-x}$	$c_3 heta^{lpha+x-1}(1- heta)^{eta+n-x-1}$
Normal/Normal	$\theta \in \mathbb{R}$	X	$N(\mu_{\scriptscriptstyle 0},\sigma_{\scriptscriptstyle 0}^{\scriptscriptstyle 2})$	$N( heta,\sigma^2)$	$N(\mu_1, \sigma_1^2)$
(fixed $\sigma^2$ )	θ	$c_1 \exp\{-\frac{1}{2\sigma_0^2}(\theta-\mu_0)^2\}$	x	$c_2 \exp\{-\frac{1}{2\sigma^2}(x-\mu)^2\}$	$c_{3} \exp\{-\frac{1}{2\sigma_{1}^{2}}(\theta-\mu_{1})^{2}\}$

Which are conjugate priors for the following pairs likelihood/prior?

- Exponential/Normal
- ② Exponential/Gamma
- Inomial/Normal

Suppose the prior has been set. Let  $x_1$  and  $x_1$  be two sets of data. Which of the following are true?

- If the likelihoods  $f(x_1|\theta)$  and  $f(x_2|\theta)$  are the same then they result in the same posterior.
- If  $x_1$  and  $x_2$  result in the same posterior then their likelihood functions are the same.
- If the likelihoods  $f(x_1|\theta)$  and  $f(x_2|\theta)$  are proportional then they result in the same posterior.
- If two likelihoods functions are proportional then they are equal.