# Lecture 3A <br> MTH6102: Bayesian Statistical Methods 

Eftychia Solea

Queen Mary University of London

2023

## Today's agenda

Today's lecture will cover

- Review Bayesian updating with continuous parameters and discrete data.
- Construct a posterior for continuous parameters and continuous data.
- Conjugate priors


## Bayesian inference

- Suppose we have data $y$ generated from $p(y \mid \theta)$ where $\theta$ is the unknown parameter.
- Start with the prior distribution $p(\theta)$ about $\theta$.
- Likelihood is $p(y \mid \theta)$.
- The resulting probability distribution $p(\theta \mid y)$ is called the posterior distribution.

$$
p(\theta \mid y)=\frac{p(\theta) p(y \mid \theta)}{p(y)} \propto p(\theta) p(y \mid \theta)
$$

Posterior distribution $\propto$ prior distribution $\times$ likelihood

- Our inferences about $\theta$ are based on this posterior distribution.


## Bayesian updating: Discrete likelihoods, continuous priors

- $\theta$ : continuous parameter with prior pdf $p(\theta)$ and range $[a, b]$.
- $x$ : random discrete data
- likelihood: $p(x \mid \theta)$

Bayesian updating table

| Hypothesis | prior prob | likelihood | Bayes numerator | posterior prob. $p(\theta \mid x) d \theta$ |
| :--- | :--- | :--- | :--- | :--- |
| $\theta$ | $p(\theta) d \theta$ | $p(x \mid \theta)$ | $p(x \mid \theta) p(\theta) d \theta$ | $\frac{p(x \mid \theta) p(\theta) d \theta}{p(x)}$ |
| Total | $\int_{a}^{b} p(\theta) d \theta=1$ |  | $p(x)=\int_{a}^{b} p(x \mid \theta) p(\theta) d \theta$ | 1 |

- The posterior density $p(\theta \mid x)$ is obtained by removing $d \theta$ from the posterior probability in the table.



## Binomial data/beta prior example

- $Y \sim \operatorname{binom}(n, q)$ with unknown binomial probability of success $q$.
- We observe $Y=k$ successes in $n$ trials. $K \sim \operatorname{binom}(n, q)$
- The binomial likelihood $p(k \mid q)$ for this problem is:

$$
P(Y=\mathcal{x} \mid q)=p(k \mid q)=\underline{\binom{n}{k} q^{k}(1-q)^{n-k}}
$$

- Convenient prior distribution for $q$ is $\operatorname{Beta}(\alpha, \beta)$ :

$$
p(q)=\frac{q^{\alpha-1}(1-q)^{\beta-1}}{B(\alpha, \beta)} \quad, q \in[0,1]
$$

## Binomial data/beta prior example



$$
p(q \mid k) \propto \underbrace{q^{k+\alpha-1}(1-q)^{n-k+\beta-1}}
$$

From this, we can recognise that the posterior distribution $p(q \mid k)$ is the $\operatorname{Beta}(k+\alpha, n-k+\beta)$ distribution.

General $\operatorname{Beta}(a, b)$ pdf: $f(x)=\frac{x^{a-1}(1-x)^{b-1}}{B(a, b)}$

We hare that the posterior density $p(q) \& 1$ is

$$
\rho(q \mid x)=c q^{x+a-1} \cdot(1-q)^{n-x+b-1}
$$

We see that $\rho(\varepsilon \mid x)$ is the beta density with parameters $x+a$ and $n-x+b$

$$
p(q|x| \sim \operatorname{beta}(x+a, n-x+e)
$$

The normalising constant $c$ mast be

$$
C=\frac{1}{B(x+a, n-x+b)}
$$

The constant $C$ maxes the density to integrate 1

## Bayesian updating table: Binomial data/beta prior

- Data $k$ generated from $\sim \operatorname{Binom}(\mathrm{n}, q)$, with $q$ unknown
- Continuous hypotheses $q$ in $[0,1]$.
- $\underline{\operatorname{Beta}(\alpha, \beta)}$ prior $p(q)$
- Binomial likelihood $p(k \mid q)$

| Hypothesis | prior prob. | likelihood | Bayes numerator | posterior prob. |
| :--- | :--- | :--- | :--- | :--- |
| $q$ | $B e t a(\alpha, \beta) d q$ | binomial $(\mathrm{n}, \mathrm{q})$ | $c q^{k+\alpha-1}(1-q)^{n-k+\beta-1} d q$ | Beta $(k+\alpha, n-k+\beta) d q$ |
| Total | 1 |  | $T=\left(\int_{0} c q^{k+\alpha-1}(1-q)^{n}\right.$ | 1 |

- The posterior density is $\operatorname{Beta}(k+\alpha, n-k+\beta)$
- Note: We don't need to compute $T$. Once we know the posterior is of the form $c q^{k+\alpha-1}(1-q)^{n-k+\beta-1}$ we have to find $c$ that makes it a proper density. In this case $c=1 / \operatorname{Beta}(k+\alpha, n-k+\beta)$


## Conjugate distributions

$$
\text { - } a+b \rightarrow \begin{aligned}
& \text { nintral namber } \\
& \text { of trids }
\end{aligned}
$$



- Binomial likelihood $k \sim \operatorname{Bin}(n, q)$
- $\operatorname{Beta}(k+\alpha, n-k+\beta)$ posterior distribution for $q \mid k$
- In this example, we have the same family of distributions for the prior and posterior distribution.
- This is known as a conjugate distribution.
- "The family of Beta distributions is conjugate to the binomial likelihood".


## Conjugate distributions

## $X \sim \operatorname{Geum}(q)$

Binomial likelihood: $p(k \mid q)=\binom{n}{k} q^{k}(1-q)^{n-k} \quad P(X=x)=q^{x}(1-q)$

Beta prior: $p(q)=\frac{q^{\alpha-1}(1-q)^{\beta-1}}{B(\alpha, \beta)}$

- Considered as functions of $q$, the prior and likelihood have the same functional form as each other (proportional to $q^{r}(1-q)^{s}$ for some $r, s)$.
- When we multiply them together, we still have the same form.
- This is what characterises conjugate distributions.

Board question
$\theta=$ probability of heads

- Suppose your prior in the bent coin example is Beta (6,8). You flip the coin 7 times, getting 2 heads and 5 tails. What is the posterior pdf $p(\theta \mid x)$ ?
(1) $\operatorname{Beta}(2,5)$
$\operatorname{Beta}(3,6)$

$$
\begin{aligned}
& \begin{array}{l}
\text { Solution } \\
a=6, \quad b=8 \\
n=7, \quad x=2, \quad n-x=5 \\
p \mid \theta(x) \sim \text { beta }(8,13)
\end{array}
\end{aligned}
$$

(2) $\operatorname{Beta}(3,6)$

Beta $(6,8)$
Beta $(8,13)$

## Board question

- A medical treatment has unknown probability $\theta$ of success.
- We assume treatment has prior $f(\theta) \sim \operatorname{Beta}(5,5)$.
(1) Suppose you test it on 10 patients and have 6 successes. Find the posterior distribution on $\theta$. Identify the type of the posterior pdf
(2) Suppose you recorded the order of the results and got SSSFFSSSFF. Find the posterior based on this new data.

Solution
$\operatorname{Prcor} \rho(\theta)=\frac{\theta^{4}(1-\theta)^{4}}{B(5,5)}, \theta \in[0,1]$
. The data $x=6$ successes in 10 patients come from a binomial model. So the likelihood is just the binomial 11 relihoud

$$
p(\sigma \mid \theta)=\binom{10}{6} \theta^{6}(1-\theta)^{4}
$$

The posterior density, $\rho(\theta(\sigma)$ is

$$
\begin{gathered}
\rho(\theta \mid 6) \propto \frac{\theta^{4}(1-\theta)^{4}}{B(5,5)} \times\binom{ 10}{6} \theta^{6}(7-\theta)^{4} \\
=C \theta^{10}(1-\theta)^{8}
\end{gathered}
$$

This is the beta $(11,9)$, so the normalising constant is just

$$
C=\frac{1}{B(11,9)}
$$

The postericu density is $\rho(\theta \mid \sigma)=\frac{\theta^{10}(1-\theta)^{8}}{B(11, q)}$.
(b) The answer is again beta $(1,9)$. The only thing that changes is the binomial coefficient which is just 1 .

## Bayesian updating: continuous priors, continuous data

We are now ready to do Bayesian updating when both the parameters and the data take continuous values.

- $\theta$ continuous parameter
- Prior pdf, $f(\theta)$

Data: continuous $x \nmid(x \mid \theta)$

- Likelihood $\sqrt{f(x \mid \theta)}$
- posterior pdf, $f(\theta \mid x)$
- Bayesian update table

| Hypothesis | prior prop | likelihood | Bayes numerator | posterior prop $f(x \mid \theta) d \theta$ |
| :--- | :--- | :--- | :--- | :--- |
| $\theta$ | $f(\theta) d \theta$ | $f(x \mid \theta)$ | $f(x \mid \theta) f(\theta) d \theta$ | $\frac{f(x \mid \theta) f(\theta) d \theta}{f(x)}$ |
| Total | 1 |  | $f^{\prime}(x)$ | 1 |

- $f(x)=\int f(x \mid \theta) f(\theta) d \theta$


## Normal example, known variance

$11 d$
$-\underbrace{y_{1}, \ldots, y_{n}} \sim \underbrace{N\left(\mu, \sigma^{2}\right) .}_{\infty}$

- It's simpler if only one parameter is unknown.
- First, consider case where only $\mu$ is unknown. $\sigma^{\partial}$ is Rnown
- Is there a conjugate prior for $\mu$ ?


## Normal example, known variance

- Observed data $\underline{y}_{1}, \ldots, y_{n} \sim \mathcal{N ( \mu , \sigma ^ { 2 } )}$ vith $\mu$ unknown and $\sigma^{2}$ known.
- Prior distribution $\mu \sim \mathcal{N}\left(\mu_{0}, \sigma_{0}^{2}\right)$.
- The posterior distribution is

$$
\mu \sim \mathcal{N}\left(\mu_{1}, \sigma_{1}^{2}\right)
$$

where

$$
\begin{gathered}
\mu_{1}=\left(\frac{\mu_{0}}{\sigma_{0}^{2}}+\frac{n \bar{y}}{\sigma^{2}}\right) /\left(\frac{1}{\sigma_{0}^{2}}+\frac{n}{\sigma^{2}}\right) \\
\sigma_{1}^{2}=1 /\left(\frac{1}{\sigma_{0}^{2}}+\frac{n}{\sigma^{2}}\right)
\end{gathered}
$$

Proof
Let $y_{1}, \ldots, y_{n} \sim N\left(\mu, \sigma^{2}\right)$ where $\sigma^{2}$ cs known. The livelihood, $p\left(y_{11}, y_{n} / \mu\right)$, is the joint density of $y_{1, \ldots}, y_{n}$ : By independence of $y_{1}, y_{n}$,

$$
\begin{aligned}
& \text { of } y_{1,}, y_{n}: \text { By independence } \\
& p\left(y_{1}, y_{n} \mid \mu\right)=\prod_{i=1}^{n} \frac{1}{\sigma \sqrt{2 \pi}} \exp \left\{-\frac{\left(y_{i}-\mu\right)^{2}}{2 \sigma^{2}}\right\} \\
& =\underbrace{\left(\frac{1}{\sigma \sqrt{2 \pi}}\right.})^{n} \exp \left\{-\sum_{i=1}^{n} \frac{\left(y_{i}-\mu\right)^{2}}{2 \sigma^{2}}\right\}
\end{aligned}
$$

$$
\alpha \exp \left\{-\sum_{i=1}^{n} \frac{\left(y_{i}-(\mu)^{2}\right.}{2 \sigma^{2}}\right\}
$$

The prov for $\mu$ is $N\left(\mu_{0}, \sigma_{0}^{2}\right)$

$$
p(\mu)=\frac{1}{\sqrt{2 \pi \sigma 0^{\circ}}} \exp \left\{-\frac{\left(\mu-\nu_{0}\right)^{2}}{2 \sigma 0^{2}}\right\}
$$

we can rewirle the likelihood as

$$
\begin{aligned}
& \text { we can rewrite the likelihood as } \\
& \left.\rho(y / \mu) \propto \exp \left\{-\sum_{i=1}^{n} \frac{\left(\bar{c}_{c}^{2}-2 y_{i} p+p^{2}\right.}{2}\right)\right\} \quad \bar{y}=\frac{2 y_{i}}{n} \\
& =\exp \left\{-\frac{1}{2 \delta^{2}}\left(\sum_{i=1}^{n} y_{c}^{2}-2 n \bar{y} \mu+n \mu^{2}\right)\right\}
\end{aligned}
$$

## Normal example, known variance

## Normal-normal Bayesian update table

- Data: $x \sim \mathcal{N}\left(\mu, \sigma^{2}\right), \sigma^{2}$ known
- Likelihood: $f(x \mid \mu)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left\{-\frac{1}{2 \sigma^{2}}(x-\mu)^{2}\right\}$.
- $\mu$ continuous with prior pdf $f(\theta) \sim \mathcal{N}\left(\mu_{0}, \sigma_{0}^{2}\right)$
- posterior $f(\mu \mid x) \sim \mathcal{N}\left(\mu_{1}, \sigma_{1}^{2}\right)$

| Hypothesis | prior prop | likelihood | Bayes numerator | posterior prop $f(x \mid \mu) d \mu$ |
| :--- | :--- | :--- | :--- | :--- |
| $\mu$ | $\frac{1}{\sqrt{2 \pi \sigma_{0}^{2}}} \exp \left\{-\frac{1}{2 \sigma_{0}^{2}}\left(\mu-\mu_{0}\right)^{2}\right\} d \mu$ | $\frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left\{-\frac{1}{2 \sigma^{2}}(x-\mu)^{2}\right\}$ | $c_{1} \exp \left\{-\frac{1}{2 \sigma_{1}^{2}}\left(\mu-\mu_{1}\right)^{2}\right\} d \mu$ | $\frac{f(x \mid \mu) f(\mu) d \mu}{f(x)}=\frac{1}{\sqrt{2 \pi \sigma_{1}^{2}}} \exp \left\{-\frac{1}{2 \sigma_{1}^{2}}\left(\mu-\mu_{1}\right)^{2}\right\} d \mu$ |
| Total | 1 |  | $f(x)=\int_{-\infty}^{\infty} c_{1} \exp \left\{-\frac{1}{2 \sigma_{1}^{2}}\left(\mu-\mu_{1}\right)^{2}\right\} d \mu$ | 1 |

## Normal example, known variance

Normal-normal updating formulas

$$
\begin{align*}
& a=\frac{1}{\sigma_{0}^{2}}, \quad b=\frac{n}{\sigma^{2}}  \tag{1}\\
& \mu_{1}=\frac{a \mu_{0}+b \bar{y}}{a+b}, \quad \sigma_{1}^{2}=\frac{1}{a+b} \tag{2}
\end{align*}
$$

- The posterior mean $\mu_{1}$ is a weighted average of the prior mean $\mu_{0}$ and sample average $\bar{y}$.
- If $n$ is large then the weight $b$ is large and $\bar{y}$ will have a strong influence on the posterior. In fact if $n \rightarrow \infty, b /(a+b) \rightarrow 1$ and $a /(a+b) \rightarrow 0$, so $\mu_{1} \rightarrow \bar{y}$.
- If $\sigma_{0}^{2}$ is small then the weight $a$ is large and $\mu_{0}$ will have a strong influence on the posterior


## Board question

- Suppose our data follows a $N(\theta, 1)$ distribution with unknown mean $\theta$.
- Suppose our prior on $\theta$ is $N(2,1)$.
- Suppose we obtain data $x=5$
- Compute the Bayesian update table and show that the posterior pdf for $\theta$ is Normal
- Find the posterior mean and the posterior variance
- Use the updating formulas (1) to find the posterior mean and posterior variance.


## Board question


prior: blue, posterior: purple, $x=5$ (data).
The posterior mean lies between the data $x=5$ and the prior mean.

## Board question


(1) Which plot is the posterior to just the first data value $x=3$ ?
(2) Which plot is the posterior to all 3 data values, $x=3, x=9$ and $x=12$ ?

## Board question

On a basketball team the free throw percentage over all players is a $N(75,36)$ distribution. In a given year individual players free throw percentage is $N(\theta, 16)$ where $\theta$ is their career average.

This season, Sophie Lee made 85 percent of her free throws.
(1) What is the posterior expected values of her career percentage $\theta$ ?

## Exponential model

- The time until failure for a type of light bulb is exponentially distributed with parameter $\lambda$.
- We observe $n$ bulbs, with failure times $t=t_{1}, \ldots, t_{n}$.
- The unknown parameter is $\lambda$.
- Can we find a conjugate family of distributions for this likelihood?


## Conjugate priors

- A prior is conjugate to a likelihood if the posterior is the same type of distribution as the prior.

|  | hypothesis | data | prior | likelihood | posterior |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Bernoulli/Beta | $\theta \in[0,1]$ | $x$ | $\operatorname{Beta}(\alpha, \beta)$ | $\operatorname{Bernoulli}(\theta)$ | $\operatorname{Beta}(\alpha+1, \beta)$ or $\operatorname{Beta}(\alpha, \beta+1)$ |
|  | $\theta$ | $x=1$ | $c_{1} \theta^{\alpha-1}(1-\theta)^{b-1}$ | $\theta$ | $c_{3} \theta^{\alpha}(1-\theta)^{\beta-1}$ |
|  | $\theta$ | $x=0$ | $c_{1} \theta^{\alpha-1}(1-\theta)^{b-1}$ | $1-\theta$ | $c_{3} \theta^{\alpha-1}(1-\theta)^{\beta}$ |
| Binomial/Beta | $\theta \in[0,1]$ | $x$ | $\operatorname{Beta}(\alpha, \beta)$ | $\operatorname{binomial}(n, \theta)$ | $\operatorname{beta}(\alpha+x, \beta+n-x)$ |
| (fixed $n)$ | $\theta$ | $x$ | $c_{1} \theta^{\alpha-1}(1-\theta)^{b-1}$ | $c_{2} \theta^{\circ}(1-\theta)^{n-x}$ | $c_{3} \theta^{\alpha+x-1}(1-\theta)^{\beta+n-x-1}$ |
| Normal/Normal | $\theta \in \mathbb{R}$ | $x$ | $N\left(\mu_{0}, \sigma_{0}^{2}\right)$ | $N\left(\theta, \sigma^{2}\right)$ | $N\left(\mu_{1}, \sigma_{1}^{2}\right)$ |
| $\left(\right.$ fixed $\left.\sigma^{2}\right)$ | $\theta$ | $c_{1} \exp \left\{-\frac{1}{2 \sigma_{0}^{2}}\left(\theta-\mu_{0}\right)^{2}\right\}$ | $x$ | $c_{2} \exp \left\{-\frac{1}{2 \sigma^{2}}(x-\mu)^{2}\right\}$ | $c_{3} \exp \left\{-\frac{1}{2 \sigma_{1}^{2}}\left(\theta-\mu_{1}\right)^{2}\right\}$ |

## Board question

Which are conjugate priors for the following pairs likelihood/prior?
(1) Exponential/Normal
(2) Exponential/Gamma
(3) Binomial/Normal

## Board question

Suppose the prior has been set. Let $x_{1}$ and $x_{1}$ be two sets of data. Which of the following are true?

- If the likelihoods $f\left(x_{1} \mid \theta\right)$ and $f\left(x_{2} \mid \theta\right)$ are the same then they result in the same posterior.
- If $x_{1}$ and $x_{2}$ result in the same posterior then their likelihood functions are the same.
- If the likelihoods $f\left(x_{1} \mid \theta\right)$ and $f\left(x_{2} \mid \theta\right)$ are proportional then they result in the same posterior.
- If two likelihoods functions are proportional then they are equal.

