# Lecture 2B <br> MTH6102: Bayesian Statistical Methods 

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## Today's agenda

Today's lecture

- Review
- Use Bayes' theorem to compute posterior pmf with discrete pmf priors.
- Use Bayes' theorem to compute posterior pfd with continuous pdf priors


## Announcements

New office hours starting from today

- Every Wednesday from 12:00-13:00, room MB-B11


## Bayesian inference

## $y \sim$ ply

- Probability model $p(y \mid \theta)$ depends on a set of parameters $\theta$.
- $\theta$ is unknown and we would like to learn about $\theta$.
- Let $y$ be the observed data, assumed to be generated by this probability model, $p(y \mid \theta) \quad y=y$ (the realsation of $y$ )
- In Bayesian statistics, we assign probabilities on both the parameters $\theta$ and data $y$


## Bayesian inference

- So we start with a probability distribution for the parameters $p(\theta)$, called the prior distribution.
- $\theta$ is either discrete or continuous random variable. Hence, $p(\theta)$ is either a pmf or pdf
- The prior is a subjective distribution, based on experimenter's belief, and is formulated before the data $y$ are seen.

Bayesian inference
－Let $y$ be the observed data．
－We then update the prior distribution（mf or pdf）to a posterior distribution（mf or pdf）for $\theta, p(\theta \mid y)$ ，using Bayes＇theorem

where the observed data enters through the likelihood $p(y \mid \theta)$ ．
－$p(y)$ is the normalising constant，which is given by

$$
p(y)=\int \frac{p\left(\theta^{\prime}\right) p\left(y \mid \theta^{\prime}\right) d \theta^{\prime} \text { or } \sum_{\theta^{\prime}} \frac{p\left(\theta^{\prime}\right)}{T_{\rho ⿰ 亻 ⿱ 丶 ⿻ 工 二 十}} p\left(y \mid \theta^{\prime}\right)}{}
$$

－$p(y)$ does not depend on $\theta$ $\theta$ is continuous $\int p(\theta / y) d \theta=1$

## What does it mean?

$$
\begin{equation*}
\left.\sim^{p(\theta \mid y}\right) \propto p(\theta) p(y \mid \theta) \tag{1}
\end{equation*}
$$

## Posterior $\propto$ prior $\times$ likelihood

- $p(y \mid \theta)$ is the likelihood and it the probability of data $y$ given the true $\theta$.
- Start with initial beliefs/information about $\theta, p(\theta)$ - this is the prior distribution formulated before the data are seen.
- Bayesian updating: Update the prior distribution using the data $y$, using (1).
- The updated prior, $p(\theta \mid y)$ is called the posterior distribution.
- We base our inferences about $\theta$ based on this posterior distribution.


## Bayesian updating with discrete data, discrete prior

- parameter $\theta$ discrete with values $\theta_{1}$ and $\theta_{2}$ and prior pmf $p(\theta)$
- Discrete data $x$
- Discrete likelihood, $p(x \mid \theta)$
- Posterior pmf: $p\left(\theta_{1} \mid x\right), p\left(\theta_{2} \mid x\right)$

$$
p\left(\theta_{1}|x|=p\left(\theta=\theta_{1}(x)\right.\right.
$$

$\rightarrow$| Hypothesis | prior | likelihood | Bayes numerator | posterior |
| :--- | :--- | :--- | :--- | :--- |
| $\theta$ | $p(\theta)$ | $p(x \mid \theta)$ | $p(x \mid \theta) p(\theta)$ | $p(\theta \mid x)$ |
| $\theta_{1}$ | $\bar{p}\left(\theta_{1}\right)$ | $p\left(x \mid \theta_{1}\right)$ | $\bar{p}\left(x \mid \theta_{1}\right) p\left(\theta_{1}\right) \cdot$ | $p\left(\theta_{1} \mid x\right)$ |
| $\theta_{2}$ | $p\left(\theta_{2}\right)$ | $p\left(x \mid \theta_{2}\right)$ | $p\left(x \mid \theta_{2}\right) p\left(\theta_{2}\right)$. | $p\left(\theta_{2} \mid x\right)$ |
| Total | 1 | NOT SUM TO 1 | $(p(x)))$ | 1 |

- Law of total probability: $p(\underline{x})=p\left(x \mid \theta_{1}\right) p\left(\theta_{1}\right)+p\left(x \mid \theta_{2}\right) p\left(\theta_{2}\right)$.
- Bayes' theorem: $p\left(\theta_{1} \mid x\right)=\frac{p\left(x \mid \theta_{1}\right) p\left(\theta_{1}\right)}{p(x)}, \quad p\left(\theta_{2} \mid x\right)=\frac{p\left(x \mid \theta_{2}\right) p\left(\theta_{2}\right)}{p(x)}$

$$
\text { posterior }=\frac{\text { likelihood } \times \text { prior }}{\text { total prob. of data }} .
$$

## Board Question: Coins

- There are three types of coins which have different probabilities of heads
- Type A coins are fair, with probability 0.5 of heads.
- Type B are bent and have probability 0.6 of heads.
- Type C are bent and have probability 0.9 of heads.

Suppose I have a drawer containing 5 coins: 2 of type A, 2 of type $B$, and 1 of type $C$. I pick a coin at random, and without showing you the coin I flip it once and get heads.

- Make a Bayesian update table and compute the posterior pmf that the chosen coin is each of the three coins.

Solution
Hypothesis. The hypothesis is the probability of heads $\theta$. The value of $\theta$ is itself random that takes three values
$\theta=0.5$ means coin is Type $A$
$\theta=0.0$ means can is Type $B$
$\theta=0.9$ means coin is Type $C$
Prov pm. Since $\theta$ is discrete, it has a prior port

$$
\begin{array}{r}
p(\theta), \theta \in\{0.5,0.6,0.9\} \\
p(0.5)=p(\theta=0.5)=\frac{2}{5} \\
p(0.6)=P(\theta=0.6)=\frac{2}{5} \\
p(0.9)=P(\theta=0.9)=\frac{1}{5}
\end{array}
$$

Data. We observe heads, $x=1$
Likelihood. The likelihood $\rho(x \mid \theta)$ is the probability of observing heads $x=1$ given $\theta$. We hare 3 li Kelihoud values

$$
\begin{gathered}
\text { Dod values } \\
p(x=1 \mid 0.5)=0.5, p(x=1(\theta=0.6)=0.6 \\
p(x=1(\theta=0.9)=0.9
\end{gathered}
$$

You con think $x$ generuled from the Bernoulli enff with unnnown probubility of heads $\underline{\theta}$ I
$x \sim \rho(x \mid \theta)$ where

$$
\rho(x \mid \theta)=\theta^{x}(1-\theta)^{1-x} \quad x=1 \text { or } x=0
$$

when $x=1 \quad p(x=1 \mid \theta)=\theta$
in ouv case, $\theta \in\{0.5,0.6,0.9\}$
Posterion pmf these are the probabilites

$$
\begin{aligned}
& p(\theta=0.5 \mid x=1) \\
& p(\theta=0.0 \mid x=1) \\
& p(\theta=0.9 \mid x=1)
\end{aligned}
$$

7 we want to
compule the postenor pmf of $\theta$

Bagesion updating table

| hy pothesu <br> $\theta$ | prow pmf <br> $\rho(\theta)$ | Li relihood <br> $\rho(x=1(\theta)$ | Bayes num <br> $\rho(x=1 \mid \theta) p(\theta)$ | posteral <br> pmf |
| :---: | :---: | :---: | :---: | :---: |
| 0.5 | $\rho(0.51=215$ | 0.5 | 0.2 | 0.3226 |
| 0.6 | $\rho(0.0)=2(5$ | 0.6 | 0.24 | 0.3871 |
| 0.9 | $\rho(0.9)=115$ | 0.9 | 0.18 | 0.2903 |
| rotal | 1 |  | $\rho(x=1)=0.62$ | $7]$ |

$$
\begin{aligned}
& \rho(x=1)=\rho\left(x=1\left(\theta_{1}\right) \rho\left(\theta_{r}\right)+\rho\left(x=1\left|\theta_{2}\right| \rho\left(\theta_{2}\right)\right.\right. \\
& \quad+\rho\left(x=1\left(\theta_{3}\right) \rho\left(\theta_{3}\right)\right. \\
& \rho\left(\theta_{1}=\theta_{1}(x=1)=\frac{0.5 \times 2(5}{0.62}=0.3226\right.
\end{aligned}
$$

- In the previous lecture, we have done Bayesian updating when we had a finite number of hypotheses or a discrete parameter $\theta$ e.g., - in the diagnostic example had 2 hypotheses (HIV + ve, HIV -ve), - in the coin example we had 3 hypothesis ( $\mathrm{A}, \mathrm{B}$ and C ).
- In this topic we will study Bayesian updating where there is a continuous range of hypotheses, i.e., $\underline{\rightarrow}$ is a continuous random variable.
- The Bayesian updating will be essentially the same, based on the Bayes' theorem
posterior $\propto$ prior $\times$ likelihood


## Examples with continuous parameters

- Suppose we have a medical treatment for a disease than can succeed or fail with probability $q$. Then $q$ is a continuous quantity between 0 and 1.
- The lifetime of a certain light bulb $T$ is modeled as an exponential distribution $\exp (\lambda)$ with unknown $\lambda$. We can assume that $\lambda$ takes any value greater than 0 .


## Baysian updating: Discrete likelihoods, continuous priors

- $\theta$ : continuous parameter with prior pdf $p(\theta)$ and range $[a, b]$.
- $x$ : random discrete data
- discrete likelihood: $p(x \mid \theta)$
- posterior pdf: $p(\theta \mid x)$
- By Bayes' theorem we update the prior pdf to a posterior pdf

$$
p(\theta \mid x)=\frac{p(x \mid \theta) p(\theta)}{p(x)}=\frac{p(x \mid \theta) p(\theta)}{\int_{a}^{b} p(x \mid \theta) p(\theta) d \theta} .
$$

- Law of total probability: $p(x)=\int_{a}^{b} p(x \mid \theta) p(\theta) d \theta$.


## Baysian updating: Discrete likelihoods, continuous priors

- $p(x)$ does not depend on $\theta$ and serves as the normalising constant so that $p(\theta \mid x)$ is a proper pdf and integrates to 1 .
- Hence, we can express Bayes' theorem in the form

$$
\begin{gathered}
p(\theta \mid x) \propto p(x \mid \theta) p(\theta) \\
\text { posterior } \propto \text { prior } \times \text { likelihood }
\end{gathered}
$$

## Bayesian inference

$$
p(\theta \mid x) \propto p(\theta) p(x \mid \theta)
$$

- $p(\theta)$ - initial beliefs/information about $\theta$, the prior pdf.
- $p(x \mid \theta)$ - the likelihood for observed data $x$ with parameters $\theta$.
- Update information about $\theta$ using the likelihood.
- The resulting pdf $p(\theta \mid x)$ is called the posterior $\operatorname{pdf}$ of $\theta$
posterior $\propto$ prior $\times$ likelihood

Bayesian updating: Discrete likelihoods, continuous priors

- Sometimes, it is better to use $\left.p^{\boldsymbol{r}} \theta\right) d \theta$ to work with probabilities instead of densities e.g the prior probability that $\theta$ is in a small interval of width $d \theta$ around 0.5 if $p(0.5) d \theta$.
- In this case, the Bayes' theorem is postereen prob


The probability that $\theta$ lies in a very small interval of aid th $d \theta$ around 0.5 is approximately $p(0.5) d \theta$
E. Solea, QMUL

Bayesian updating: Discrete likelihoods, continuous priors

- $\theta$ : continuous parameter with prior pdf $p(\theta)$ and range $[a, b]$.
- $x$ : random discrete data
- likelihood: $p(x \mid \theta)$

Bayesian updating table

$\rightarrow$| Hypothesis | prior prob | likelihood | Bayes numerator | posterior prob. $p(\theta) x) d \theta$ |
| :--- | :--- | :--- | :--- | :--- |
| $\theta$ | $p(\theta) d \theta$ | $p(x \mid \theta)$ | $0(x \mid \theta) p(\theta) d \theta$ | $\frac{p(x \mid \theta) p(\theta) d \theta}{p(x)}$ |
| Total | $\int_{a}^{b} p(\theta) d \theta=1$ |  | $\left(\rho(x)=\int_{a}^{b} p(x \mid \theta) p(\theta) d \theta\right.$ | 1 |

- The posterior density $p(\theta \mid x)$ is obtained by removing $d \theta$ from the posterior probability in the table.

$$
\varphi\left(\theta(x) d \theta=\frac{\varphi(x|\theta| \rho|\theta| d \theta}{\rho(x)}\right.
$$

The posterior density is $p(\theta) \pi)$

## Bayesian updating: Discrete likelihoods, continuous priors

## Example: Binomial data, Beta prior

- A biased coin has probability of heads $q$ which is unknown.
- We toss the coin $n$ times and observe $k$ heads (This is my data $x=k$ ).
- The binomial likelihood for this problem:

$$
\underbrace{p(k \mid q)}=\binom{n}{k} q^{k}(1-q)^{n-k} \rightarrow \begin{gathered}
\text { binomial } \\
\text { imf }
\end{gathered}
$$

- For Bayesian inference, we need to specify a prior distribution for $q$.
- $q$ is a continuous quantity between 0 and 1 .
- What is a possible probability distribution for $q$ (or family of distributions)?


## Bayesian updating: Discrete likelihoods, continuous priors

## Example: Binomial data, Beta prior

- The family of Beta distributions seems a natural choice for a prior distribution for $q$, since it describes continuous random variables with support on $[0,1]$.
- If $q \sim \operatorname{Beta}(\alpha, \beta)$, its probability density function is

$$
f(q)=\underbrace{\frac{q^{\alpha-1}(1-q)^{\beta-1}}{B(\alpha, \beta)}}, 0 \leq q \leq 1,
$$

where $B$ is the Beta function and $\alpha$ and $\beta$ are parameters,

$$
\begin{gathered}
B(\alpha, \beta)=\int_{0}^{1} x^{\alpha-1}(1-x)^{\beta-1} d x \\
\int^{B(\alpha, \beta)=} \frac{\Gamma(\alpha) \Gamma(\beta)}{\Gamma(\alpha+\beta)}
\end{gathered}
$$

Beta distributions
Beta $(1,1)$ is the uniform distribution on [or]

- Probability density functions.
- If $q \sim \operatorname{Beta}(\alpha, \beta)$ $\frac{\mathcal{E ( q )}=\frac{\alpha}{\alpha+\beta}}{\text { Trial means }}$




## Bayesian updating: Discrete likelihoods, continuous priors

## Example: Binomial data, Beta prior

- Bayesian updating: posterior $\propto$ prior $\times$ likelihood $x=K$
- The posterior distribution $p(q \mid \sqrt{*})$ is proportional to

$$
p(q \mid k) \propto \underbrace{q^{k+\alpha-1}(1-q)^{n-k+\beta-1}}
$$

- We recognise this to have the form of a beta distribution, so the posterior is a beta distribution, beta $(k+\alpha, n-k+\beta)$.
- Hence, the normalising constant must be $1 / B(k+\alpha, n-k+\beta)$.

Proof: The posterior density, $p(\varepsilon \mid x)$, is $p(q|x| \propto p(x \mid q) p(q)=$ Bayes numerator

$$
\begin{aligned}
& =C_{1}\binom{n}{x} q^{x}(1-q)^{n-x} \frac{r(a+b)}{F(a) F(b)} q^{a-1}(1-q)^{b-1} \\
& =C_{2} q^{x+a-1}(1-q)^{n-x+b-1}
\end{aligned}
$$

We recognise this to hare the same form with the beta pdf. Hence, we can find the value of $\mathrm{C}_{2}$ that mates $p(q \mid x)$ a proper density (integrates to 1)

$$
\begin{aligned}
& \text { We want } \\
& B(x+a, n-x+e) \cdot C_{2} \int_{0}^{1} \frac{q^{x+a-1}(1-q)^{n-x+b-1}}{B(x+a, n-x+b)} \\
& \text { Beta }(x+a,=1,-x+e) \text { density }
\end{aligned}
$$

$$
\begin{array}{r}
\Rightarrow B\left(x+a_{1} n-x+b\right) \cdot C_{2}=1 \\
\quad \Rightarrow C_{2}=\frac{1}{B\left(x+a_{1} n-x+b\right)}
\end{array}
$$

The posterior density $p(q \mid x)$ is Beta $(x+a, n-x+b)$

$$
p(q \mid x)=\frac{q^{x+a-1}(1-q)^{n-x+b-1}}{B(x+a, n-x+e)}, \quad q \in[0,1]
$$

You don't need to compute

$$
\begin{aligned}
& P(x=x)=\rho_{0} \mid \rho(\varepsilon) d q \\
& \rho(q \mid x)=\frac{\rho(x \mid q) \rho(\varepsilon)}{\rho(x=x)}=\frac{\text { Bases numerator }}{\text { Total porto. of heads }}
\end{aligned}
$$

## Bayesian updating: Discrete likelihoods, continuous priors

- The actual, normalized pdf is

$$
p(q \mid k)=\frac{q^{k+\alpha-1}(1-q)^{n-k+\beta-1}}{B(k+\alpha, n-k+\beta)}
$$

the pdf of a $\operatorname{Beta}(k+\alpha, n-k+\beta)$ r.v. (Remember: the random variable is $q$ and $k$ is fixed).

- Bayesian updating: We update the prior $\operatorname{Beta}(\alpha, \beta)$ to posterior $\underbrace{\operatorname{Beta}(k+\alpha, n-k+\beta)}$. CFov binomiol data)


## Bayesian updating table: Discrete likelihoods, continuous priors

## Example: Binomial data, Beta prior

- $Y \sim \operatorname{Binom}(\mathrm{n}, q)$, with $q$ unknown
- Continuous hypotheses $q$ in $[0,1]$.

- Prior $p(q)$
- Likelihood $p(y \mid q)$

| Hypothesis | prior prob. | likelihood | Bayes numerator | posterior prob. |
| :--- | :--- | :--- | :--- | :--- |
| $q$ | $\operatorname{Beta}(\alpha, \beta) d q$ | binomial $(\mathrm{n}, \mathrm{q})$ | $c q^{k+\alpha-1}(1-q)^{n-k+\beta-1} d q$ | $\operatorname{Beta}(k+\alpha, n-k+\beta) d q$ |
| Total | 1 |  | $T=\int_{0}^{1} c q^{k+\alpha-1}(1-q)^{n-k+\beta-1} d q$ | 1 |

- The posterior density is $\operatorname{Beta}(k+\alpha, n-k+\beta)$
- Note: We don't need to compute $T$. Once we know the posterior is of the form $c q^{k+\alpha-1}(1-q)^{n-k+\beta-1}$ we have to find $c$ that makes it a proper density. In this case $c=1 / \operatorname{Beta}(k+\alpha, n-k+\beta)$


## Unknown parameters and prior parameters

## Remarks:

- We need to distinguish between the parameters we are estimating, which we generally have denoted by $\theta$ and the parameters for the prior distribution (s).
- In this binomial example, $q$ is uncertain: we have prior and posterior distributions for $q$.
- The parameters of the prior distribution, here $\alpha$ and $\beta$, are taken as fixed.

$$
p(q) \sim \operatorname{beta}(a, b) \text { where } \begin{aligned}
& a_{1} b \text { are fixed } \\
& \text { and known }
\end{aligned}
$$

$x \sim \operatorname{binom}(n, q) q$ is unknown

## Board question: bent coin

- Bent coin with unknown probability $\theta$ of heads
- Prior: $p(\theta)=2 \theta$ on $[0,1]$
- Data: toss and get heads
- Compute the Bayesian update table.
- Find the posterior pdf to this data.

Solution
Hypothesis is $\theta$, the pwbability of heads, $\theta \in[0,1]$
Prior pdf. $\rho(\theta)=2 \theta, \theta \in[0,1]$
Data: $x=1$, in this care we observe heads urelihood The Ii Kelinoud, $p(x=1 \mid \theta)$ is the probability of observing heads given the twi $\theta$. Then

$$
p(x=1 \mid \theta)=\theta
$$

Postenor pdf The posteriow density, $p(\theta \mid x=1)$ is $\varphi(\theta \mid x=1) \propto \varphi(x=1|\theta| \times \rho(\theta)$ - Bones numerator

$$
\begin{aligned}
& =c_{1} \theta \times 2 \theta \\
& =c_{2} \theta^{2}
\end{aligned}
$$

We want to find $C_{2}$ such thee $\int_{0}^{1} p(\theta \mid x=1) d \theta=1$
So $C_{2} \int_{0}^{0} \theta^{\theta} d \theta=1 \Leftrightarrow C_{2}\left[\frac{\theta^{3}}{3}\right]_{0}^{1}=1 \Rightarrow C_{2}=3$
The posterior $\rho d f$ of $\theta$ given the data, $x=1$, is

$$
p(\theta \mid x=1)=\underbrace{3 \theta^{2}}_{\operatorname{Beta}(3,1)}, \quad \theta \in[0,1]
$$

Bayesion Table

| hypothesis <br> $\theta$ | $2 \theta d \theta$ | $\theta$ | $c_{2} \theta^{2} d \theta$ | $3 \theta^{2} d \theta$ |
| :---: | :---: | :---: | :---: | :---: |
| Total prob. | 1 |  | II Kelihoud | Bages numer. |

## Posterior mean

In Bayesian how would you choose a particular value of $q$ ?

- A natural estimate for $q$ is the mean of the posterior distribution $p(q \mid k)$, called the posterior mean.
- For the binomial case with $\operatorname{Beta}(\alpha, \beta)$ prior, the posterior mean is

$$
\left.\hat{q}_{\mathrm{B}}=E(q\rceil k\right)=\frac{k+\alpha}{n+\alpha+\beta}
$$

- The prior distribution has mean $\alpha /(\alpha+\beta)$ which would be our best estimate of $q$ without having observed the data.
- Ignoring the prior, we would estimate $q$ using the maximum likelihood estimate (MLE)

$$
\hat{q}=\frac{k}{n}
$$

- The Bayes' estimate $\hat{q}_{B}$ combines all of this information.
if $q \sim \operatorname{Beta}(a, b)$, then $\mathbb{E}(\varepsilon)=\frac{a}{a+b}$ prior mean

$$
\text { If } \begin{aligned}
& q \mid x \sim \operatorname{Beta}(\underset{\sim}{x+a, n-x+b)}, I E(q \mid x)=\frac{x+a}{\not x+a+n-x+b} \\
&=\frac{x+a}{n+a+b} \\
& \underbrace{}_{\text {posterior mean }}
\end{aligned}
$$

- Note that we can rewrite $\hat{q}_{b}$ as

$$
\hat{q}_{\mathrm{B}}=\frac{n}{n+\alpha+\beta}\left(\frac{k}{n}\right)+\frac{\alpha+\beta}{n+\alpha+\beta} \frac{\alpha}{\alpha+\beta}
$$

- Thus $\hat{q}_{\mathrm{B}}$ is a linear combination of the prior mean and the MLE, with the weights being determined by $n, \alpha$ and $\beta$
$\hat{q}_{B \rightarrow B}$ stands for posterior meon/Bayesion estimate / Bayes estimate
- One important prior is called flat prior or uniform prior.
- A flat prior assumes that every hypothesis is equally probable. formef betcol
- For example if $q$ has range $[0,1]$, the $p(q)=1$ is a flat prior.
- E.g. a uniform distribution on $[0,1]$ is $\underbrace{\operatorname{Beta}(1,1)}_{r}$

$$
a=1, b=1
$$

- So, posterior distribution is $\operatorname{Beta}(\underbrace{k+1, n-k+1})$
- Posterior mean: $E(q \mid k)=\frac{k+1}{n+2}$


## Board question

- Bent coin with unknown probability $\theta$ of heads
- Flat prior: $p(\theta)=1$ on $[0,1]$
- Data: toss 27 times and get 15 heads and 12 tails.
- Compute the Bayesian update table.
- Give the integral for the normalising factor but do not compute it. Call its value $T$ and give the posterior pdf in terms of $T$.


## Board question

- A medical treatment has unknown probability $\theta$ of success.
- We assume treatment has prior $f(\theta) \sim \operatorname{Beta}(5,5)$.
(1) Suppose you test it on 10 patients and have 6 successes. Find the posterior distribution on $\theta$. Identify the type of the posterior pdf
(2) Suppose you recorded the order of the results and got SSSFFSSSFF. Find the posterior based on this new data.

