

Main Examination period 2023 – January – Semester A

## MTH6102: Bayesian Statistical Methods

**Duration: 2 hours**

The exam is intended to be completed within **2 hours**. However, you will have a period of **4 hours** to complete the exam and submit your solutions.

**You should attempt ALL questions. Marks available are shown next to the questions.**

All work should be **handwritten** and should **include your student number**. Only one attempt is allowed – **once you have submitted your work, it is final**.

In completing this assessment:

- You may use books and notes.
- You may use calculators and computers, but you must show your working for any calculations you do.
- You may use the Internet as a resource, but not to ask for the solution to an exam question or to copy any solution you find.
- You must not seek or obtain help from anyone else.

When you have finished:

- scan your work, convert it to a **single PDF file**, and submit this file using the tool below the link to the exam;
- e-mail a copy to **maths@qmul.ac.uk** with your student number and the module code in the subject line;

**Examiners: J. Griffin, D. Stark**

**Question 1 [24 marks].**

Suppose that we have data  $y = (y_1, \dots, y_n)$ . Each data-point is assumed to be generated by a distribution with the following probability density function:

$$p(y_i | \psi) = 2\psi y_i \exp(-\psi y_i^2), \quad y_i \geq 0, \quad i = 1, \dots, n.$$

The unknown parameter is  $\psi$ , with  $\psi > 0$ .

- (a) Write down the likelihood for  $\psi$  given  $y$ . Find an expression for the maximum likelihood estimate (MLE)  $\hat{\psi}$ . [6]
- (b) A Gamma( $\alpha, \beta$ ) distribution is chosen as the prior distribution for  $\psi$ . Derive the resulting posterior distribution for  $\psi$  given  $y$ . [6]
- (c) Show that the posterior mean for  $\psi$  is always in between the prior mean and the MLE for this example. [5]
- (d) The data are  $y = (2, 6, 5, 4, C + 1)$ , where  $C$  is the last digit of your ID number, with  $n = 5$ . The prior distribution is Gamma(2, 2).
- (i) What is the MLE  $\hat{\psi}$ ? [3]
- (ii) What is the posterior distribution for  $\psi$ ? Based on this posterior distribution, calculate a point estimate for  $\psi$ . [4]

**Question 2 [19 marks].**

The data  $y = (y_1, \dots, y_n)$  is a sample from a normal distribution with unknown mean  $\mu$  and known standard deviation  $\sigma = 2$ . The prior distribution for  $\mu$  is normal  $N(\mu_0, \sigma_0^2)$ . The posterior distribution is  $\mu | y \sim N(\mu_1, \sigma_1^2)$ , where

$$\mu_1 = \left( \frac{\mu_0}{\sigma_0^2} + \frac{n\bar{y}}{\sigma^2} \right) / \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right), \quad \sigma_1^2 = 1 / \left( \frac{1}{\sigma_0^2} + \frac{n}{\sigma^2} \right), \quad \text{and } \bar{y} \text{ is the sample mean.}$$

- (a) As the prior distribution becomes less informative, what value does the posterior mean for  $\mu$  approach? As the prior distribution becomes more informative, what value does the posterior mean for  $\mu$  approach? [4]
- (b) Suppose that we take  $\mu_0 = 0$ , and we want the prior probability  $P(|\mu| \leq A + 20)$  to be 0.9, where  $A$  is the third-to-last digit of your ID number. What value for  $\sigma_0$  should we choose? [4]

Let the sample mean be  $B + 1$ , where  $B$  is the second-to-last digit of your ID number, and the sample size be  $n = 40$ . Use the prior distribution found in part (b).

- (c) What is the posterior distribution for  $\mu$ ,  $p(\mu | y)$ ? What is the posterior median for  $\mu$ ? [4]
- (d) Let  $x$  be a future data-point from the same  $N(\mu, \sigma^2)$  distribution. Find the posterior predictive mean and variance of  $x$ . [7]

**Question 3 [26 marks].**

The dataset  $y = (y_1, \dots, y_n)$  is a sample from a Poisson distribution with parameter  $\lambda$ . A  $\text{Gamma}(\alpha, \beta)$  prior distribution is assigned to  $\lambda$ . Apart from part (c), the answers do not need any numerical calculations. In the following R code, the data  $y$  is denoted by  $y$  in the code, and  $\alpha$  and  $\beta$  are the prior parameters.

```
alpha = 3
beta = 3
a = sum(y) + alpha
b = length(y) + beta
pgamma(2, shape=a, rate=b)
qgamma(c(0.5, 0.025, 0.975), shape=a, rate=b)
```

- (a) In statistical terms, what will the last line of code output? [5]
- (b) What will the line which starts with `pgamma` output? [2]
- (c) Let  $B$  and  $C$  be the second-to-last and last digits of your ID number, respectively. Take the sample size  $n = B + 15$ , and  $\sum_{i=1}^n y_i = C + 30$ . What are the posterior mean and standard deviation for  $\lambda$ ? [5]

The R code below follows on from the code above.

```
v = rgamma(5000, shape=a, rate=b)
w = rpois(length(v), lambda=v)
mean(w==0)
```

- (d) When this code has run, what will  $v$  contain? What will  $w$  contain? [6]
- (e) What quantity will the last line of code output (in statistical terms)? [3]
- (f) State one advantage of using a prior distribution which is conjugate to the likelihood. [2]
- (g) Suppose that we assumed some other prior distribution instead of a gamma distribution. What method could we use to make inferences based on the resulting posterior distribution for  $\lambda$ ? [3]

**Question 4 [16 marks].**

The observed data is  $y = (y_1, \dots, y_n)$ , a sample from a geometric distribution with parameter  $q$ . The prior distribution for  $q$  is uniform on the interval  $[0, 1]$ . Suppose that  $y_1 = \dots = y_n = 0$ . Take  $n = 10 + A$ , where  $A$  is the third-to-last digit of your ID number.

- (a) What is the normalized posterior probability density function for  $q$ ? [5]

Suppose now that we want to compare two models. Model  $M_1$  assumes that the data follow a geometric distribution with  $q$  known to be  $q_0 = 0.8$ . Model  $M_2$  is the model and prior distribution described above.

- (b) Find the Bayes factor  $B_{12}$  for comparing the two models. [6]
- (c) We assign prior probabilities of  $1/2$  that each model is the true model. Find the posterior probability that  $M_1$  is the true model. [3]
- (d) State a drawback of using Bayes factors and posterior probabilities to compare models. [2]

**Question 5 [15 marks].**

The observed data  $y = \{y_{ij}, i = 1, \dots, n, j = 1, \dots, m_i\}$  are the average results in an exam for school  $j$  within county  $i$ . The following hierarchical model is considered reasonable:

$$y_{ij} \sim \text{Normal}(\mu_i, \sigma_S^2), j = 1, \dots, m_i$$

$$\mu_i \sim \text{Normal}(\mu_C, \sigma_C^2), i = 1, \dots, n.$$

where  $\mu_C$ ,  $\sigma_S$  and  $\sigma_C$  are unknown parameters which are each assigned a prior distribution. Suppose that we have generated a sample of size  $M$  from the joint posterior distribution  $p(\mu_C, \sigma_S, \sigma_C, \mu_1, \dots, \mu_n | y)$ .

- (a) Explain how to use the posterior sample to estimate the following:
- (i) the posterior mean for  $\mu_C$ ;
  - (ii) a 95% credible interval for  $\sigma_S / \sigma_C$ ;
  - (iii) the posterior probability that  $\mu_1 < \mu_2$ . [7]
- (b) Explain how to generate a sample from the posterior predictive distribution of the result for a school not in our dataset, in each of the following two cases:
- (i) if the county containing the school is in our dataset;
  - (ii) or if the county is not in our dataset. [8]

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**End of Paper – An appendix of 1 page follows.**

## Appendix: common distributions

For each distribution,  $x$  is the random quantity and the other symbols are parameters.

### Discrete distributions

Distribution	Probability mass function	Range of parameters and variates	Mean	Variance
Binomial	$\binom{n}{x} q^x (1-q)^{n-x}$	$0 \leq q \leq 1$ $x = 0, 1, \dots, n$	$nq$	$nq(1-q)$
Poisson	$\frac{\lambda^x e^{-\lambda}}{x!}$	$\lambda > 0$ $x = 0, 1, 2, \dots$	$\lambda$	$\lambda$
Geometric	$q(1-q)^x$	$0 < q \leq 1$ $x = 0, 1, 2, \dots$	$\frac{(1-q)}{q}$	$\frac{(1-q)}{q^2}$
Negative binomial	$\binom{r+x-1}{x} q^r (1-q)^x$	$0 < q \leq 1, r > 0$ $x = 0, 1, 2, \dots$	$\frac{r(1-q)}{q}$	$\frac{r(1-q)}{q^2}$

### Continuous distributions

Distribution	Probability density function	Range of parameters and variates	Mean	Variance
Uniform	$\frac{1}{b-a}$	$-\infty < a < b < \infty$ $a < x < b$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Normal $N(\mu, \sigma^2)$	$\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$	$-\infty < \mu < \infty, \sigma > 0$ $-\infty < x < \infty$	$\mu$	$\sigma^2$

The 95th and 97.5th percentiles of the standard  $N(0, 1)$  distribution are 1.64 and 1.96, respectively.

Exponential	$\lambda e^{-\lambda x}$	$\lambda > 0$ $x > 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Gamma	$\frac{\beta^\alpha x^{\alpha-1} e^{-\beta x}}{\Gamma(\alpha)}$	$\alpha > 0, \beta > 0$ $x > 0$	$\frac{\alpha}{\beta}$	$\frac{\alpha}{\beta^2}$
Beta	$\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$	$\alpha > 0, \beta > 0$ $0 < x < 1$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$

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**End of Appendix.**